# A STUDY OF INTERCOHORT CHANGE IN WOMEN'S WORK PATTERNS AND EARNINGS 

M. Anne Hill and June E. O'Neill

Center for the Study of Business and Government Baruch College

FINAL REPORT
December 1990

This Project was funded by the U.S. Department of Labor, Bureau of Labor Statistics under Grant Number E-9-J-8-0092. Opinions stated in this document do not necessarily represent the official position or policy of the U.S. Department of Labor. We gratefully acknowledge the excellent data preparation and programming assistance provided by Hengzhong Liu, Partha Deb, Chun-Yong Yang, and T. Jithendranathan. Additional funding for this research has been provided by The Rockefeller Foundation.

# A STUDY OF INTERCOHORT CHANGE IN WOMEN'S WORK PATTERNS AND EARNINGS 

M. Anne Hill and June E. O'Neill

Center for the Study of Business and Government Baruch College

FINAL REPORT
December 1990

This Project was funded by the U.S. Department of Labor, Bureau of Labor Statistics under Grant Number E-9-J-8-0092. Opinions stated in this document do not necessarily represent the official position or policy of the U.S. Department of Labor. We gratefully acknowledge the excellent data preparation and programming assistance provided by Hengzhong Liu, Partha Deb, Chun-Yong Yang, and T. Jithendranathan. Additional funding for this research has been provided by The Rockefeller Foundation.

## TABLE OF CONTENTS

EXECUTIVE SUMMARY
I. INTRODUCTION ..... 1
II. CHANGES IN LIFE-CYCLE LABOR FORCE PATTERNS, SCHOOLING, AND FERTILITY ..... 8
A. Background ..... 8
B. Description of the Data ..... 9
C. Descriptive Analysis ..... 13

1. Work Participation ..... 13
2. Cumulative Years of Work Experience ..... 19
3. Years of Schooling ..... 26
4. Changes in Marital status and Fertility ..... 26
D. Summary ..... 35
III. DETERMINANTS OF THE DURATION OF WORK AND NON-WORK SPELLS ..... 36
A. Multiple Spell Hazard Rate Model ..... 36
5. Methodology ..... 38
6. Description of the Data for the Multiple Spell Model ..... 40
7. Empirical Analysis ..... 44
IV. DETERMINANTS OF THE PROPORTION OF YEARS WORKED OVER THE LIFETIME ..... 56
A. Empirical Results ..... 56
B. Accounting for the Increase in Lifetime Participation ..... 63
C. Summary ..... 68
V. IS THE FEMALE EXPERIENCE-WAGE PROFILE GROWING STEEPER? ..... 70
A. Cohort/Age/Race Specific Regressions ..... 70
B. Regressions Using Pooled Data ..... 73
REFERENCES ..... 78
APPENDIX A
APPENDIX

After remaining virtually constant during the post-World War II period, the ratio of women's earnings to men's increased sharply during the 1980's, rising from 59.7 percent in 1979 to 68.5 percent in 1989. The failure of the overall wage gap to narrow during the 1950-1980 period has been something of a puzzle. The labor force participation of women had escalated over the entire post World War II period, the women's movement had blossomed, and barriers to women's entry into many professions and occupations appeared to have eroded. Yet, overall women's relative earnings did not rise through the 1960s and 1970s.

Explanations for the apparent paradox have pointed out that during the post War period of rapid increases in the proportion of women who work, the experience level and other work related attributes of the average employed women did not increase, as less experienced women joined the ranks of the employed. The failure of the wage gap to narrow before 1980, therefore, can be partly explained by the failure of women's lifetime work experience to rise as new labor force entrants lowered the work experience of the average working woman.

Quite the opposite phenomenon may underlie the increase in women's relative earnings in the 1980 s. Indeed, one explanation for this recent development is that there has been an intercohort increase in the continuity of women's labor force participation resulting in a greater accumulation of work experience. The fact that the recent gains in women's relative wages have been larger among younger women suggests that intercohort trends in human capital acquisition are important and that more recent cohorts of women have gained relative to men in terms of marketable skills.

Yet because women are heterogeneous with respect to lifetime experience, changes in accumulated work experience cannot be measured using standard cross-sectional data. Longitudinal or retrospective data on life-cycle experience are crucial for studying issues related to women's labor supply, human capital formation, and earnings. This research utilizes data from the three continuing panels of the National Longitudinal Surveys (NLS) -- the mature women, the young women, and the youth cohort -- to measure accumulated years of work experience and to examine changes in life-cycle work patterns across successive cohorts of women born between. 1923 and 1964.

This study has investigated how these successive cohorts of women have changed with respect to their accumulation of work-related skills, in terms of level of schooling, career orientation, and attachment to the labor force. We consider how the nature of entry into and exit from the labor force changed across cohorts and how the response of women's labor force participation decisions to life-cycle events (e.g., marriage, the birth of a child, divorce) may have changed. Intercohort changes in women's returns to work experience, schooling, and other human capital investments are also considered. This research has yielded important insights into the nature and determinants of the work patterns and earnings of American women.

Our comparison of human capital and demographic characteristics across these seven cohorts of women has illuminated the dramatic changes in labor market experience and its correlates. Labor force participation, whether
measured at a point in time or over the lifetime has increased markedly for white women, with black women experiencing slight increases or declines. For working and nonworking women combined, the cumulative years during which an individual has worked at least aix months has risen, although the average level of experience of employed women has grown more slowly or has actually declined. While some of this slower growth can be attributed to the lower levels of experience held by new entrants and the rapid increase in the number of new entrants (as signalled by the rise in survey week participation rates), rising levels of schooling have also diminished the number of post-schooling years within which women (at a fixed age) could have worked.

Along with rising levels of investment in education, these cohorts of women have experienced dramatic demographic changes. A larger proportion of each cohort remains unmarried and more women continue to be childless. Moreover, the number of children ever born among these women has declined sharply, with women in the earlier cohorts bearing three to four children and more recent cohorts giving birth to two to three children.

We have examined the determinants of work experience using two models. For the NLSY and the NLS young women, we were able to eatimate the multiple spell hazard rate model of work status transitions. Moreover, we have examined, for all cohorts, ordinary least squares regressions of the proportion of possible years worked as of a given age.

Both the duration model and the lifetime participation model yield strikingly similar results. During the period 1967 to 1987 , the length of work spells and the proportion of the lifetime worked have increased. These changes are strongly related to rising levels of schooling, delayed childbearing and reductions in fertility, and transformations in marriage patterns. Results from the duration models imply that labor supply responses are becoming increasingly sensitive to schooling and prior work experience (especially among black women). And for white women, much of the intercohort change in lifetime participation appears to result from dramatic fertility declines. Yet the estimates from pooled models for lifetime participation indicate that, holding constant the effects of the independent variables, there remains a strong, statistically significant effect of the passage of time. Even with identical characteristics, women from more recent cohorts are spending a higher proportion of their time at work in the market.

Finally, we have estimated wage models for these cohorts of women to investigate whether or not experience-wage profiles have grown steeper over time. Our results provide evidence that work-related investments have increased from cohort to cohort among white women, although not necessarily for all cohorts of black women. And while we cannot determine from our analysis the extent to which women or their employers are responsible for the increased levels of investment, the former pattern of flat age-earnings profiles for women -- the dead-end job syndrome -- finally appears to have been overcome, which bodes well for future narrowing of the gender wage gap.

> A STUDY OF INTERCOHORT CHANGE IN WOMEN'S WORK PATTERNS AND EARNINGS
> M. Anne Hill and June E. O'Neill
> Center for the Study of Business and Government Baruch College

## I. INTRODUCTION

After remaining virtually constant during the post-World War II period, the ratio of women's earnings to men's increased sharply during the 1980's, rising from 59.7 percent in 1979 to 68.5 percent in 1989 (based on the annual earnings of full-time year-round workers aged 16 years and over). As Figure 1 indicates, relative wage gains were experienced both by black women and white women in the 1980s. However, the wage gap between black women and men began to narrow several decades ago and it has been smaller than the gap between white women and men since the 1960 s.

The failure of the overall wage gap to narrow during the $1950-1980$ period has been something of a puzzle. The labor force participation of women had escalated over the entire post World War II period, the women's movement had blossomed, and barriers to women's entry into many professions and occupations appeared to have eroded. Yet, overall women's relative earnings did not rise through the 1960 s and 1970 .

Explanations for the apparent paradox have pointed out that during the post War period of rapid increases in the proportion of women who work, the experience level and other work related attributes of the average employed women did not increase, as less experienced women joined the ranks of the employed (Goldin, 1989; O'Neill, 1985; Smith and Ward, 1989). It has been demonstrated in a number of studies that a primary factor underlying the gender gap in wages is the gender differential in the amount and continuity of lifetime work experience (e.g., Mincer and Polachek, 1974, 1978; Corcoran and

Figure 1
Trends in Women's Relative Wages


Duncan, 1979). The failure of the wage gap to narrow before 1980 , therefore, can be partly explained by the failure of women's lifetime work experience to rise as new labor force entrants lowered the work experience of the average working woman.

Quite the opposite phenomenon may underlie the increase in women's relative earnings in the $1980 s$. Indeed, one explanation for this recent development is that there has been an intercohort increase in the continuity of women's labor force participation resulting in a greater accumulation of work experience. The fact that the recent gaing in women's relative wages have been larger among younger women suggests that intercohort trends in human capital acquisition are important and that more recent cohorts of women have gained relative to men in terms of marketable skills.

Figure 2 illustrates lifetime patterns of labor force participation for successive cohorts of women born between 1926-30 and 1961-65. These cohorts roughly correspond to the panels of women included in the National Longitudinal Surveys which are the focus of this study. For each cohort the figure shows the change in labor force participation rates at different ages in the life cycle. Two observations stand out. One is that there has been a large rise in participation from one cohort to the next at all stages of the life cycle. Second, the shape of the cohort profiles have changed. Among earlier birth cohorts, participation rates decline between ages 20 to 24 and 25-29, reflecting lower levels of labor force activity during ages when young children are present. However, the cohort profiles then rise, particularly between ages 30 to 34 and 45 to 49 . Thus, the labor force participation rate of the cohort born 1941-1945 was 45 percent at ages 25 to 29 but rose to 74 percent when they reached ages 45 to 49.

Figure 2
Cohort Participation Rates


When the profile line rises steeply, as it does for the cohort born 1941-45 it is evident that a large proportion of those in the labor force at ages 45 to 49 could not have worked continuously over much of their adult years. By contrast, the cohorts born $1951-55$ and thereafter, have much flatter cohort profiles.- Their -labor force participation is initially high at ages 20 to 24 and continues to rise without experiencing the characteristic dip or the steep rebound of the earlier cohorts. The flatter profiles of these recent cohorts, combined with their high levels of participation hint that a larger proportion of women currently in their thirties have accumulated more years of work experience than was the case among earlier cohorts. Since the younger cohorts have yet to be observed over the life-cycle, the extent to which their participation will eventually decline or will continue to rise throughout the life-cycle remains to be seen.

The fact that the recent gains in women's relative wages have been largest among younger women, and decline systematically with age, guggests that intercohort trends in human capital acquisition are important and that more recent cohorts of women have gained relative to men in terms of marketable skills. ${ }^{1}$ -

It is important to recognize that changes in accumulated work experience cannot be constructed from the standard cross-sectional data. Because women are heterogeneous with respect to lifetime work experience, longitudinal or retrospective data on life-cycle experience are crucial for studying issues

[^0]related to women's labor supply, human capital formation, and earnings. This research utilizes data from the three continuing panels of the National Longitudinal Surveys (NLS) -- the mature women, the young women, and the youth cohort -- to measure accumulated years of work experience and to examine changes in life-cycle work patterns across successive cohorts of women born between 1923 and 1964.

In this report, we provide detailed descriptive analysis of how marital and fertility patterns, schooling attainment, labor force participation, and work attachment have changed from cohort to cohort. We estimate the factors associated with entry into and exit from the labor force for a given cohort and consider how the responses to these factors compare across cohorts. We also analyze the determinants of lifetime labor force participation and the wages.

We address the following broad questions:

1. How have successive cohorts of women changed with respect to their accumulation of work-related akills? In particular, how have life-cycle patterns of labor force participation changed? And what have been the intercohort changes in schooling levels and in career orientation? 2. What are the factors associated with entry and exit from the labor force for a given cohort and how do the responses to these factors compare across cohorts? Factors potentially influencing the duration of labor force spells include schooling, births, changes in marital status, and changes in labor market conditions.
2. How have returns to work experience, schooling and other human capital invegtments changed between successive cohorts of women?

The following section describes the data in greater detail and by
comparing changes across cohorts in labor force participation, schooling, marriage and fertility, provides a background for our empirical analysis of the determinants of these changes. The third and fourth sections focus on the estimation of economic models of lifetime work decisions, with section three describing the results-for-a dynamic work speli model, and section four reporting the results of regressions for the proportion of possible years worked. The fifth section provides estimates of wage modela to ascertain the extent to which the underlying intercohort changes in accumulation of human capital have influenced wages.

## A. Background

In this section we summarize our findings on intercohort changes in women's patterns of labor force participation at a moment in time and over the life cycle."..The basic-economic framework for analyzing women's labor gupply derives from the family's allocation of its members time among market work, home work, and leisure. ${ }^{2}$ The extent of market work depends on the expected remuneration from market work relative to the "ghadow wage" from home production. The theory, therefore, predicts that economic growth would increase women's participation in the market provided that the substitution effect of rising relative market wages dominates the income effect of rising family income on leisure (and the demand for home produced goods). Evidence on the dominance of the gubstitution effect has been found in a number of cross-sectional studies, and some time series studies, using U.S. data as well as data from other countries (e.g., Mincer, 1962; Cain, 1966; Hill, 1983; 0'Neil1, 1981, and the articles in Layard and Mincer, 1985). 3

It is unlikely that so fundamental a change as the shift of women'g work activity from the home to the market would be a simple story of response to rising wage levels; and it is not. Changes in the labor force participation of women are closely intertwined with changes in fertility and marital stability, and it has proven difficult to separate cause and effect. Some decline in fertility may have been exogenous (for example, due to developments in contraceptives), but to a large extent fertility must surely be a decision

[^1]jointly determined with labor force participation. The dramatic increase in the divorce rate in the United States also may have some exogenous component associated with liberalization in divorce laws. However, marital instability is also likely to increase as a result of women's rising labor force participation.(which reduces the gain to marriage); and labor force participation may in turn increase in response to rising marital instability (and concomitant uncertainty regarding income).

We therefore examine intercohort changes in important variables that are associated with labor force participation -- schooling, marital status, and fertility. Our analysis is based on three unique panel surveys. The following section describes these data.
B. Description of the Data

The data for this analysis have been drawn from two panels of the National Longitudinal Surveys of Labor Market Experience (NLS) - - the mature women and the young women -- and from the panel of young women in the National Longitudinal Survey of Youth (NLSY).

The NLS panel of mature women were born in the years 1923 to 1937 and were first surveyed in 1967, when they were ages 30 to 44 . Personal and telephone interviews were conducted at regular intervals over the years (most recently in 1988). The NLS young women, born 1944-53 were interviewed initially in 1968 when they were 14 to 24 years of age. Of the initial sample of 5,083 mature women, an estimated 3,346 remained in the sample in 1986 . Among young women, 3,720 of the original 5,159 remained by 1985. The sample of NLSY women numbers 6,282. This panel was born in 1959 to 1965 and was first interviewed in 1979 at ages 14 to 21. The NLS panels provide a superior source of data for this analysis since
the longitudinal nature of the survey permits direct observations on work experience, earnings, and other related variables. Unfortunately, the surveys of mature women and young women began skipping years in the 1970\%, so that a full bequence of annual observations on all variables is not available. However, through the use of retraspective questions; it has been possible to complete much information for the missing years on work experience, fertility, and other important variables. Our methodology for accomplishing this will be described in detail below.

For this analysis, we divide the samples into five-year age cohorts and view these cohorts over time. Combining these groups yields seven cohorts observed at five-year intervals, as outlined in Table 1. As indicated in this table, the NLS Mature Women are observed in 1967, 1972, 1977, and 1982 (and 1968, 1973, 1978, and 1983 for the Young Women). The surveys during these pivotal years were taken using detailed in-person interviews, and they consequently yield greater detail regarding demographic characteristics. Moreover, questions for these years often gleaned retroapective information that can be combined with the panel data to create a complete time series for important variables.

While these surveys include detailed retrospective components for estimating characteristics such as schooling, fertility, and marital status (inter alia), surveys for the mature women, young women, and NLSY differ markedly in the quality of information available regarding years of work experience.

In the initial survey questionnaire for the mature women (in 1967), respondents were asked to report the number of years since completing schooling during which they had worked at least six months. For subsequent

Table 1
NLS and NLSY Cohorts by Age, Survey Year and Cohort Number

*NLSY cohorts 6 and 7 overlap.
years, we define a "work year" as one in which the respondent worked 26 weeks or more (or at least half of the possible weeks if the "weeks worked" question pertains to more than 52 weeks). The sum of retrospective years and observed years worked yields the cumulative years of experience. This measure, defined as years since school completion, can be constructed for both the mature women and the NLS young women. However, both surveys skipped two years during the 1970s for which the NLS failed to obtain weeks worked. The NLS young women's survey in 1978 included a retrospective question asking how many of the past five years had the respondent worked at least six months. For the young women, then, we can infer for most respondents whether or not the two missing years were work years by comparing the response to this retrospective question to the prospective information combined with all employment information available (especially tenure and employment history details). Unfortunately, the detailed 1977 survey of the NLS mature women failed to include a similar retrospective "past five years" question. The available information on items such as tenure and job history rosters provided complete coverage for only a portion of the sample. Consequently, for the mature women, we estimate the number of years worked during the five-year interval 1973 to 1977 by evaluating the proportion of the three known years worked (i.e., $0,1 / 3,2 / 3$, or 1) and assigning that proportion to the remaining two years. The computer programs used to generate these data are available from the authors. The women from the NLSY are observed annually in the panel. A retrospective question fills in years worked since age 18 for the older members of the sample. Since the NLSY were not asked years worked since school completion, we calculate for them the number of years since age 18 during which they worked at leagt 2.6 weeks. Unfortunately, we can construct a
comparable measure for only one other age cohort, NLs young women who were 15 to 19 in 1968 (for whom 1967 data are reported). Because the year left achool is not fixed and many women return to school at a later time (which is increasingly so), the measure of years worked since age 18 is likely to be more accurately and consistently reported across.individuals and cohorts than the more subjective measure of years worked since leaving school.
C. Descriptive Analysis

## 1. Fork Participation

Participation in labor market activity can be described in several ways. We begin by examining labor force participation rates during the survey week, which are the rates traditionally used by the Bureau of Labor statigtics to measure labor market activity. The labor force participation rate shows the percentage of women in the population who are either employed or unemployed (as opposed to those who are out of the labor force and are not seeking work).

Table 2 illustrates these survey week participation rates by five-year age group, race and survey year. When we compare participation rates of a given age group in different years -- e.g. the group ages 30-34 in 1967, 1978, and 1983 -- we are making an intercohort comparison since we are comparing the rate of cohorts born at different times and reaching the designated age in the years specified. ${ }^{4}$ Among white women, participation rates have grown steadily over time at each age level. For example, among white women 30 to 34,44 percent were participants in 1967. This rate increased to 61.4 percent in 1978 and further to 71.2 percent in 1983.

[^2]Table 2
Survey Week Participation Rates, Weighted

| Age | Survey Year |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \\ & 1979 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \end{aligned}$ | 1987 |
| Black |  |  |  |  |  |
| 15-19 | 32.6 |  | 44.7 |  |  |
| 20-24 | 61.9 | 60.1 |  | 65.4 |  |
| 25-29 | - | 63.2 | 64.0 |  | 66.1 |
| 30-34 | 62.3 |  | 67.2 | 75.2 | - |
| 35-39 | 69.2 | 61.8 |  | 76.7 |  |
| 40-44 | 67.8 | 61.9 | 69.5 |  | -- |
| 45-49 |  | 66.0 | 64.9 | 67.8 |  |
| 50-54 |  |  | 63.9 | 63.8 |  |
| NonBlack |  | - ... |  | - | - . - |
| 15-19 | 37.7 |  | 55.2 |  |  |
| 20-24 | 58.7 | 65.8 |  | 73.2 |  |
| 25-29 |  | 56.1 | 67.6 |  | 67.2 |
| 30-34 | 44.0 |  | 61.4 | 71.2 |  |
| 35-39 | 47.6 | 53.7 |  | 73.2 |  |
| 40-44 | 51.0 | 56.1 | 62.2 |  |  |
| 45-49 |  | 57.1 | 61.2 | 67.3 |  |
| 50-54 |  |  | 58.4 | 59.4 |  |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

The picture for black women differs. In the earliest period, black women's participation rates were in excess of 60 percent at ages 20 and over and were, for the most part, considerably above the rates of white women. However, during this twenty-year period, black women's participation rates actually declined in the early 1970 s and then increased, but more slowly than white women's. Since the 1970 , white women's participation has been well above black women's at younger ages (15 to 24) when the higher fertility of black women is a factor. At ages above 25 , participation rates no longer differ between black and white women to any significant degree.

This measure of participation corresponds to activity reported within one week. If there is considerable labor force turnover, the survey week participation rate can underestimate the proportion of women who work at aome point during a year, but will overestimate the proportion working a full year. Our interest here, however, is participation over the life cycle. We firgt estimate lifetime participation in terms of the proportion of years worked since leaving school, defined as the number of years during which the respondent works 26 or more weeks divided by the total number of years eince being enrolled full-time in school. Tables 3 and 4 present such lifetime measures first for all women, and then for the group of women who were employed in the survey week.

Among all white women, the proportion of years worked has increased systematically, with rises corresponding to those in the survey week rates. In fact, the survey week participation rates are quite similar to the lifetime

Table 3
Proportion of Years Worked Since Leaving School, or Since Age 18, All Women, Weighted

|  | Since Leaving School |  |  |  | Since | Age 18 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \end{aligned}$ | 1978 | 1987 |
| Black |  |  |  |  |  |  |
| 25-29 | - -- - | 0.463 | 0.564 |  | 0.486 | 0.516 |
| 30-34 | 0.556 |  | 0.557 | 0.574 |  |  |
| 35-39 | 0.560 | 0.550 |  | 0.577 |  |  |
| 40-44 | 0.599 | 0.587 | 0.579 |  |  |  |
| 45-49 |  | 0.658 | 0.618 | 0.603 |  |  |
| 50-54 |  |  | 0.618 | 0.621 |  | - |
| NonBlack |  |  |  |  |  |  |
| 25-29 |  | 0.523 | 0.663 |  | 0.583 | 0.673 |
| 30-34 | 0.463 |  | - 0.559 | 0.621 |  |  |
| 35-39 | 0.453 | 0.466 |  | 0.568 |  | - |
| 40-44 | 0.454 | 0.462 | 0.494 |  |  |  |
| 45-49 | $\cdots$ | 0.461 | 0.488 | 0.519 |  | - |
| 50-54 |  |  | 0.476 | 0.506 |  |  |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

Table 4
Proportion of Years Worked Since Leaving School, or Since Age 18, Employed Women, Weighted

|  | Since Leaving School |  |  |  | Since | Age 18 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \end{aligned}$ | 1978 | 1987 |
| Black |  |  |  |  | - |  |
| 25-29 |  | 0.626 | 0.741 |  | 0.640 | 0.634 |
| 30-34 | 0.672 |  | 0.718 | 0.732 |  |  |
| 35-39 | 0.643 | 0.670 |  | 0.704 |  |  |
| 40-44 | 0.730 | 0.685 | 0.698 |  |  |  |
| 45-49 | . | 0.723 | 0.706 | 0.738 |  |  |
| 50-54 |  |  | 0.748 | 0.711 |  |  |
| NonBlack |  |  |  |  |  |  |
| 25-29 |  | 0.699 | 0.822 |  | 0.703 | 0.753 |
| 30-34 | 0.645 |  | 0.730 | 0.755 |  |  |
| 35-39 | 0.610 | 0.612 |  | 0.693 | - |  |
| 40-44 | 0.594 | 0.600 | 0.638 |  |  |  |
| 45-49 |  | 0.576 | 0.622 | 0.651 |  |  |
| 50-54 |  |  | 0.612 | 0.635 | .- - | - |

Source: NLS Surveys of Mature Women, Young Women, and Youth.
measures, although they diverge somewhat in recent years. ${ }^{5}$ For example, the survey week participation rate for white women 35 to 39 was 47.6 percent in 1967 and the lifetime rate for the same group was then 45.3 percent. The lifetime rate has increased more slowly than the survey week rate, rising to 56.8 percent for women 35 to. 39 years old in 1983, while the corresponding participation rate was 73.2 percent. Among black women, lifetime participation rates have changed in ways gimilar to the survey week rate: falling slightly among the older groups and increasing siightly among the younger groups.

Table 4 presents these lifetime participation rates for the subgroup of women who were employed during the survey week. If women were perfectly homogeneous so that employed women had the same prior work experience as those out of the labor force the rate for the employed would be the game as the rate for all women. However, this is not the case. Across all race and age groups, the rates for employed women exceed those for all women, implying that women who are currently working are much more likely to have been working (26 weeks or more) in prior years. For example, currently employed white women ages 30 to 34 in 1983 had worked 75.5 percent of all possible years since completing school, while the average for all women was 62.1 percent. These rates have increased among all white women, though the effect is largest at younger ages, and among younger black women. This implies that the average working woman in the late 1970 s and early $1980 s$ possessed higher levels of
${ }^{5}$ The lifetime participation rate for all women (employed and nonemployed) would be equivalent to the average. of the survey week participation rates in each year over the cohort's years since leaving school if the lifetime rate were defined as the proportion of years worked one week or more. Our estimate is more restrictive since it counts a year of work if 26 or more weeks were worked during the year.
prior work experience than her counterpart in the last 1960s. It is also noteworthy that at ages 40 and over black working women have considerably more prior work experience than white women although that difference has been narrowing.

We consider also the distribution of the proportion of years worked, displayed in Tables 5 and 6 for black and white women, respectively. Considering black women at ages 40 and older, rising proportions have worked more than half of all possible years since completing schooling, although the proportion of women working all possible years has fallen among these groups. For example, in 1967, 13.3 percent of all black women 40 to 44 had worked each year since completing school. This share fell to 6.4 percent by 1977. Among the younger groups of black women, there appears to be greater heterogeneity, and this diversity is increasing somewhat. For example, while the proportion of black women 30 to 34 who worked some (but not all) years remained gtable, the proportion working all years rose from 9.3 percent in 1967 to 11.4 percent in 1983. Also, the proportion working no years rose from 6.0 percent to 8.1 percent during the same period.

Table 6 indicates that the patterns in these proportions for white women are reasonably consistent, with a rising proportion of women at all ages working more than half of all possible years, and generally, falling proportions of women who have never worked. It is interesting to note that, especially among younger black women, there are higher proportions who have never worked than among white women.

## 2. Cumulative Years of Work Experience

We turn to consider the effect that the dramatic increases in labor

Table 5
Percent Distribution of Black Women by Proportion of Years Worked

|  | Since Leaving School |  | Since Age | 18 |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | 1967 | 1972 | 1977 | 1982 |  |  |
| Age | 1968 | 1973 | 1978 | 1983 | 1978 | 1987 |


| 25-29 |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 0 |  | 24.7 | 14.0 |  | 14.4 | 13.4 |
| .01-. 49 |  | 26.0 | 26.5 |  | 33.1 | 30.9 |
| .50-. 99 |  | 38.2 | 37.4 |  | 47.0 | 45.9 |
| 1 |  | 11.1 | 22.1 |  | 5.6 | 9.8 |
| 30-34 |  |  |  |  |  |  |
| 0 | 6.0 |  | 13.0 | 8.1 |  |  |
| .01-. 49 | 32.1 |  | 26.0 | 28.7 |  |  |
| .50-. 99 | 52.6 |  | 48.6 | 51.8 |  |  |
| 1 | 9.3 |  | 12.4 | 11.4 |  |  |
| 35-39 |  |  |  |  |  |  |
| 0 | 4.9 | 3.2 |  | 6.8 |  |  |
| .01-. 49 | 36.9 | 35.4 |  | 31.4 |  |  |
| .50-. 99 | 48.3 | 55.7 |  | 54.5 |  |  |
| 1 | 9.9 | 5.8 |  | 7.4 |  |  |
| 40-44 |  |  |  |  |  |  |
| 0 | 2.1 | 2.3 | 1.9 |  |  |  |
| .01-. 49 | 35.7 | 35.3 | 35.4 |  |  |  |
| .50-. 99 | 48.8 | 55.6 | 56.2 |  |  |  |
| 1 | 13.3 | 6.8 | 6.4 |  |  |  |
| 45-49 |  |  |  |  |  |  |
| 0 |  | 1.5 | 1.2 | 1.9 |  |  |
| . 01-. 49 |  | -32.5 | 32.5 | 32.0 |  |  |
| $.50-.99$ |  |  |  | 61.2 |  |  |
| $1$ |  | 7.7 | 7.2 | 4.9 |  |  |
| 50-54 |  |  |  |  |  |  |
| 0 |  |  | 0.7 | 1.2 |  |  |
| .01-. 49 |  |  | 31.3 | 29.2 |  |  |
| .50-. 99 |  |  | 60.0 | 61.5 |  |  |
| 1 |  |  | 8.0 | 8.2 |  |  |

Source: ÑLS Surveys of Mature women, Young women, ana Youth.

Table 6
Percent Distribution of Nonblack Women by Proportion of Years Worked

|  |  | Since Leaving School |  |  | Since Age 18 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Age | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \end{aligned}$ | $\begin{aligned} & 1.982 \\ & 1983 \end{aligned}$ | 1978 | 1987 |
| 25-29 |  |  |  |  |  |  |
| 0 |  | 17.5 | 6 |  | 5.3 | 4.5 |
| .01-. 49 |  | 25.4 | 22 |  | 29.2 | 21.1 |
| . 50-. 99 |  | 40.0 | 41 |  | 54.7 | 51.2 |
| 1 |  | 17.1 | 30. |  | 10.8 | 23.2 |

30-34

| 0 | 5.2 |
| :---: | ---: |
| $.01-.49$ | 52.7 |
| .50 .99 | 31.8 |
| 1 | -.10 .3 |


| 6.9 | 2.5 |
| ---: | ---: |
| 33.6 | 28.6 |
| 46.5 | 56.6 |
| 13.0 | 12.3 |

35-39

| 0 | 3.7 | 3.0 | 4.5 |
| :---: | ---: | ---: | ---: |
| $.01-.49$ | 56.1 | 53.0 | 35.8 |
| $.50-.99$ | 30.6 | 37.2 | 51.2 |
| 1 | 9.6 | 6.8 | 8.5 |

40-44

2.8
3.0
55.3
35.9
5.8
1.4
$.01-.49$
$.50-.99$
1
56.8
33.8
6.6
51.1
40.7

45-49
$.01-.49$
$.50-.99$
1
1.7
1.5
0.4
57.0
53.0
46.7

50-54

$$
\begin{gathered}
0 \\
.01-.49 \\
.50-.99 \\
1
\end{gathered}
$$

$$
\begin{array}{rr}
0.9 & 0.6 \\
55.0 & 48.4 \\
40.7 & 47.2 \\
3.6 & 3.7
\end{array}
$$

Source: NLS Surveys of Mature Women, Young Women, and Youth.
force participation have had on cumulative years of work experience. Table 7 diaplays for all women the average cumulative years Bince leaving school during which the respondents worked 26 weeks or more by cohort and race. Among all black women the average level of experience has either fallen or remained steady through this period. For example, women 35 to 39 possessed, on average, 11.2 years of experience in 1967 and only 9.7 in 1983. All white women, however, have gained experience, with increases of roughly one year on average among all groups. And while older black women exhibit higher levels of experience on average, years of work experience among younger white women has begun to exceed that of their black counterparts.

Table 8 restricts the sample to women who were employed in the survey week and considers how levels of cumulative experience among employed women has changed. Levels of work experience among employed women clearly exceed the population average, and experience gains have been modest among employed white women, with declines experienced by some groups. The patterns among currently employed black women reflect the changes in the overall population. In 1982 (or 1983), employed black women in most age groups posseas fewer years of experience than did gimilax women in earlier years.

Table 9 displays the cumulative years of experience since age 18 , which we calculate for the NLSY and for the youngest group of the NLS young women. We also present data for men from the NLSY. Among black women 25 to 29, the average years of experience since age 18 increased for the entire population by one-tenth of one year between 1978 and 1983 , from 4.4 to 4.5 , but actually fell for the average employed black woman. For white women, average experience increased both for the average employed woman and the average woman overall, although by a smaller amount for employed women.

Table 7
Cumulative Years of Experience Since School Completion, All Women, Weighted

| Survey Year |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Age | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & \text { 1977* } \\ & 1978 \\ & 1979 \end{aligned}$ | $\begin{aligned} & 1982 * \\ & 1983 \end{aligned}$ |  |
| Black |  |  |  |  |  |
| 25-29 |  | 3.6 | 4.0 |  |  |
| 30-34 | 8.6 |  | 6.4 | 7.3 |  |
| 35-39 | 11.2 | 11.1 |  | 9.7 |  |
| 40-44 | 15.0 | 14.5 | 14.4 |  |  |
| 45-49 |  | 18.2 | 18.2 | 18.1 |  |
| 50-54 | - -- |  | 21.5 | 21.5 |  |
| NonBlack |  | * | - .. |  |  |
| 25-29 |  | 3.7 | 4.4 |  |  |
| 30-34 | 6.6 |  | 6.4 | 7.7 |  |
| 35-39 | 8.8 | 9.0 |  | 9.5 |  |
| 40-44 | 11.2 | 11.3 | 12.0 |  |  |
| 45-49 |  | 13.7 | 14.4 | 15.1 |  |
| 50-54 | - |  | 16.5 | 17.4 | $\cdots$ |

*Mature women's years of experience calculated using ratio
method.
Source: NLS Surveys of Mature Women, Young Women and Youth.

Table 8
Cumulative Years of Experience Since School Completion, Employed Women, Weighted

| Age | Survey Year |  |  | $\begin{aligned} & 1982 * \\ & 1983 \end{aligned}$ | $\cdots$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 * \\ & 1978 \\ & 1979 \end{aligned}$ |  |  |
| Black |  |  |  |  |  |
| 25-29 |  | 4.9 | 5.1 | . | -- |
| 30-34 | 10.2 |  | 8.1 | - 9.3 |  |
| 35-39 | 12.6 | 13.2 |  | 11.7 |  |
| 40-44 | 18.1 | 16.6 | 17.1 |  |  |
| 45-49 |  | 21.7 | 20.6 | 21.9 | - |
| 50-54 |  |  | 25.8 | 24.4 | - |
| NonBlack | $\cdots$ | - |  | --. | F- |
| 25-29 |  | 4.7 | 5.2 |  |  |
| 30-34 | 9.2 |  | 8.2 | 9.3 |  |
| 35-39 | 11.9 | 11.9 |  | - 11.5 |  |
| 40-44 | 14.6 | 14.6 | 15.4 |  |  |
| 45-49 |  | 17.0 | 18.2 | 18.9 |  |
| 50-54 |  |  | 21.0 | 21.9 |  |

*Mature women's years of experience calculated using ratio method.

Source: NLS Surveys of Mature Women, Young Women, and Youth.

Table 9
Cumulative Years of Experience Since Age 18, by Employment Status (Weighted)

|  |  | Women |  | Men |
| :---: | :---: | :---: | :---: | :---: |
| Age | 1978 | 1983 | 1987 | 1987 |
| Black |  |  |  |  |
| 25-29 |  |  |  |  |
| All | 4.4 |  | 4.5 | 5.8 |
| Employed | 5.7 |  | 5.5 | 6.2 |
| Out of the Labor Force | 2.2 |  | 2.6 | 4.0 |
| 30-34 |  |  |  |  |
| All |  | 7.7 |  |  |
| Employed |  | 9.9 |  | - |
| Out of the Labor Force |  | 3.5 |  |  |
| NonBlack |  |  |  |  |
| 25-29 |  |  |  |  |
| All | 5.2 |  | 5.9 | 6.9 |
| Employed | 6.3 |  | 6.5 | 7.2 |
| Out of the Labor Force | 3.8 |  | 4.5 | 5.4 |
| 30-34 |  |  |  |  |
| All |  | 8.5 |  |  |
| Employed Out of the Labor Force |  | 10.3 5.7 |  |  |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

This comparison of years of experience between all women and currently employed women clearly indicates that while labor force participation has been rising rapidly and the cumulative years of work experience have increased for the population of all women, experience for the average employed woman has undergone glower-growth, and in some cases has even declined.

Yet some of these apparent declines in work experience since schooling (and as of a given age) may reflect the rising levels of education attained by these groups of women, and consequently, fewer years within which to work at any fixed age. We turn now to consider changes in schooling attainment.

## 3. Years of Schooling

As Table 10 indicates, level of women's schooling has increased considerably during this period. Rising levels of educational attainment have been achieved by black and white women in all age groups, with gains as large as two years for some groups. Perhaps more notable than the overall growth is the educational gains that black women have made relative to white women of similar ages. The schooling difference has narrowed from 1.0 to 1.5 years in 1967-68 to 0.2 to 0.9 years among the most recent groups of women under 35 . Table 11 reports schooling attainment levels for currently employed women. Working women are, on average, better educated than in the population at large, with one-third to one-half years more schooling attained than the average for all women.

## 4. Changes in Marital Status and Fertility

Rising investments in human capital and labor force attachment can be expected to affect (and be influenced by) marital status and fertility. Table 12 exhibits the proportion of the cohort that has never married by given

Table 10
Years of Schooling Completed by All Women, Weighted

| Age | Survey Year |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \\ & 1979 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \end{aligned}$ | 1987 |
| Black |  |  |  |  |  |
| 15-19 | 9.95 |  | 9.56 |  | - |
| 20-24 | 11.37 | 11.65 |  | 12.35 |  |
| 25-29 |  | 11.57 | 12.15 |  | 12.72 |
| 30-34 | 10.27 |  | 11.67 | 12.38 |  |
| 35-39 | 10.29 | 10.54 |  | 11.69 |  |
| 40-44 | 9.77 | 10.44 | 11.12 |  |  |
| 45-49 |  | 10.13 | 10.91 | 11.07 |  |
| 50-54 | . |  | 10.52 | 10.72 |  |
| NonBlack |  |  |  |  |  |
| 15-19 | 10.48 |  | 9.70 |  |  |
| 20-24 | 12.46 | 12.63 |  | 12.53 |  |
| 25-29 | - | $-12.76$ | 13.11 |  | 13.10 |
| 30-34 | 11.79 |  | 12.99 | 13.27 |  |
| 35-39 | 11.60 | 12.08 |  | 13.21 |  |
| 40-44 | 11.17 | 11.88 | 12.20 |  |  |
| 45-49 |  | 11.37 | 12.05 | 12.32 | . - |
| 50-54 |  |  | 11.53 | 12.16 |  |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

Table 11
Years of Schooling Completed by Employed Women, Weighted

| Age | Survey Year |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \end{aligned}$ | 1987 |
|  |  |  | 1979 |  |  |
| Black |  |  |  |  |  |
| 15-19 | 10.42 |  | 10.70 |  |  |
| 20-24 | 11.89 | 11.92 |  | 12.81 |  |
| 25-29 |  | 12.11 | 12.72 |  | 13.21 |
| 30-34 | 10.18 |  | 12.41 | 12.99 |  |
| 35-39 | 10.89 | 11.08 |  | 12.38 |  |
| 40-44 | 10.38 | 11.12 | 11.62 |  |  |
| 45-49 |  | 10.57 | 11.60 | 11.51 | - - - |
| 50-54 |  |  | 11.27 | 11.38 |  |
| NonBlack |  |  |  |  |  |
| 15-19 | 11.34 |  | 10.25 | 12.88 | - |
| 20-24 | 12.64 | 12.71 |  |  |  |
| 25-29 |  | 13.31 | 13.54 |  | 13.42 |
| 30-34 | - 11.83 |  | 13.30 | 13.54 |  |
| 35-39 | 11.57 | 12.22 |  | 13.53 |  |
| 40-44 | 11.28 | 11.99 | 12.50 |  |  |
| 45-49 |  | 11.60 | 12.42 | 12.68 |  |
| 50-54 |  |  | 11.96 | 12.46 |  |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

Table 12
Proportion of Cohort Never Married, Weighted


Source: NLS Surveys of Mature Women, Young Women, and Youth.
ages. At all ages younger than 40 , white women have experienced declining marriage rates. In 1967 , only 5 percent of white women had never married by age 30 to 34. This proportion has more than doubled (to 11 percent) in 1983. While 12 percent of women 25 to 29 had never married in 1973, 28 percent remained never married as of this age in 1987. And among black women, the proportion who never married rose even more sharply during this period, e.g., from 25 percent to 53 percent among black women 25 to 29 and from 12 percent to 29 percent among women 30 to 34 .

Among black and white women under 40 , the proportion of women who remained childless also rose during this period, as illustrated in Table 13. While only 7.5 percent of black women 35 to 39 had no children in 1967, 14.7 percent had no children by these ages in 1983. Comparable declines in childbearing were experienced by white women. And while working women are more likely than average to be childless, these proportions have also risen, as illustrated in Table 14. An increase in the proportion of women who have no children at a given age can signal either delayed childbearing or reduced total fertility, or both. Tables 15 and 16 illustrate the numbers of children ever born to all women and employed women, respectively. These data indicate striking reductions in total fertility for women younger than 40. Average declines in children born are highest among black women. For example, by ages 35 to 39 , women in 1967 had borne 4.6 children yet only 2.6 in 1983. Black women 30 to 34 had already given birth to almost 4 children in 1967, and only to about 2 in 1983. Throughout this period, working women have borne fewer children than average (data for them are depicted in Table 16). However, difference between the number of children ever borne by working women and all women has narrowed recently, as women with children have increased their work

Percentage of Women with No Children By Age and Race, All Women, Weighted

Survey Year

|  | 1967 | 1972 | 1977 | 1982 | 1987 |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | 1968 | 1973 | 1978 | 1983 |  |  |
| Age |  |  | 1979 |  |  |  |

Black

| $20-24$ |  | 0.433 |  | 0.504 |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $25-29$ |  | 0.210 | 0.204 |  | 0.303 |
| $30-34$ | 0.058 |  | 0.160 | 0.134 |  |
| $35-39$ | 0.075 | 0.053 |  | 0.147 |  |
| $40-44$ | 0.182 | 0.064 | 0.053 |  |  |
| $45-49$ |  | 0.175 | 0.064 | 0.059 |  |
| $50-54$ |  |  | 0.179 | 0.064 |  |
| NonBlack |  |  |  |  |  |


| $20-24$ |  | 0.701 |  | 0.737 | 0.483 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $25-29$ |  | 0.327 | 0.429 |  | 0.267 |
| $30-34$ | 0.116 |  | 0.201 | 0.156 |  |
| $35-39$ | 0.102 | 0.093 |  |  |  |
| $40-44$ | 0.117 | 0.094 | 0.092 |  |  |
| $45-49$ |  | 0.114 | 0.093 | 0.086 |  |
| $50-54$ |  |  | 0.114 | 0.090 |  |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

Table 14
Percentage of Employed Women with No Children, Weighted

| Age | Survey Year |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \\ & 1979 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \end{aligned}$ | 1987 |
| Black |  |  |  |  |  |
| 20-24 |  | 0.446 |  | 0.646 |  |
| 25-29 |  | 0.277 | 0.280 |  | 0.384 |
| 30-34 | 0.078 |  | 0.189 | 0.163 |  |
| 35-39 | 0.087 | 0.075 |  | 0.144 |  |
| 40-44 | 0.222 | 0.083 | 0.058 |  |  |
| 45-49 |  | 0.179 | 0.086 | '0.077 |  |
| 50-54 |  |  | 0.225 | 0.083 |  |
| NonBlack |  |  |  |  |  |
| 20-24 |  | 0.824 |  | 0.847 |  |
| 25-29 |  | 0.539 | 0.610 |  | 0.614 |
| 30-34 | 0.241 |  | 0.320 | 0.379 |  |
| 35-39 | 0.179 | 0.139 |  | 0.209 |  |
| 40-44 | 0.178 | 0.129 | 0.125 |  |  |
| 45-49 |  | 0.142 | 0.109 | 0.094 |  |
| 50-54 |  |  | 0.134 | 0.102 |  |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

Table 15
Number of Children Ever Born, By Age and Race, All Women, Weighted


Source: NLS Surveys of Mature Women, Young Women, and Youth.

> Table 16
> Number of Children Ever Born, By Age and Race Employed Women, Weighted

| Age | Survey Year |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \\ & 1979 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \end{aligned}$ | "1987 |
| Black |  |  |  |  |  |
| 20-24 |  | 0.863 |  | 0.500 |  |
| 25-29 |  | 1.690 | 1.292 |  | 1.085 |
| 30-34 | 3.342 |  | 2.227 | 1.725 |  |
| 35-39 | 4.205 | 3.832 |  | 2.464 |  |
| 40-44 | 2.604 | 4.239 | 3.792 |  |  |
| 45-49 |  | 2.860 | 4.024 | 3.863 |  |
| 50-54 |  |  | 2.680 | 4.343 |  |
| NonBlack |  |  |  |  |  |
| 20-24 | -. | 0.240 |  | 0.208 |  |
| 25-29 |  | 0.791 | 0.625 |  | 0.625 |
| 30-34 | 2.188 |  | 1.379 | 1.171 |  |
| 35-39 | 2.449 | 2.558 |  | 1.760 |  |
| 40-44 | 2.421 | 2.718 | 2.769 |  |  |
| 45-49 |  | 2.646 | 2.899 | 2.879 |  |
| 50-54 |  | $\cdots$ | 2.673 | 2.950 |  |

Source: NLS Surveys of Mature Women, Young Women, and Youth.
participation.

## D. Summary

Our comparison of human capital and demographic characterigtics across these seven cohorts of women has illuminated the dramatic changes in labor market experience and-ita correlates. . Labor force participation, whether measured at a point in time or over the lifetime has increased markedly for white women, with black women experiencing slight increases or declines. For working and nonworking women combined, the cumulative years during which an individual has worked at least six months has risen, although the average level of experience of employed women has grown more slowly or has actually declined. While some of this slower growth can be attributed to the lower levels of experience held by new entrants and the rapid increase in the number of new entrants (as signalled by the rise in survey week participation rates), rising levels of schooling have also diminished the number of post-schooling yeara within which women (at a fixed age) could have worked.

Along with rising levels of investment in education, these cohorts of women have experienced dramatic demographic changes. A larger proportion of each cohort remains unmarried and more women continue to be childless. Moreover, the number of children ever born among these women has declined sharply, with women in the earlier cohorts bearing three to four children and more recent cohorts giving birth to two to three children.

## III. DETERMINANTS OF THE DURATION OF WORK AND MON-WORK SPELLS

We have observed dramatic changes in fertility, marital status, and schooling, along with equally striking intercohort increases in work attachment, whether measured by contemporaneous participation or lifetime experience....We employ two methodologies to help sort out the influences of demographic changes on women's life-cycle labor supply behavior. In this section we estimate, for NLS Young Women and the NLSY, a multiple spell hazard rate model of exits from work and non-work spells. Unfortunately, the requirements of the multiple spell model outstripped the data available from the Mature Women's sample. However, in our second approach we estimate for all cohorts the determinants of the proportion of the lifetime worked as of a given age. These results are described in section IV which follows.
A. Multiple Spell Hazard Rate Model

This research builds on recent work using longitudinal data to analyze labor supply behavior and the implications of labor force patterns for the wage gap. ${ }^{6}$ Notable among them is Donohue (1987) who uses NLS data to estimate the duration of the first post-schooling job for samples of young men and women from two four-year periods, 1968-1971 and 1979-1982, reflecting different cohorts. His results indicate that for the more recent cohort the early labor force behavior of young women appears virtually identical to that of young men.

Donohue reports that in the early period, young women appear to leave their first jobs at considerably higher rates than do young men. However, in

[^3]the later period, the early labor force behavior for young women appears virtually identical to that of young men. Especially, the distribution by job tenure and occupation, and the behavioral estimates of his hazard rate model differ little by gender. The closing of the "first-job tenure gap" appears associated with narrowing the wage gap. -For his sample, Donohue reporta that the female-male wage ratio for white high school graduates of 77 percent for 1968-1971 rises to 89 percent by 1979-82. Donohue hypothesizes that the increasing gimilarity in behavior derives both from greater labor force attachment of young women as well as from increases in the age at marriage and first birth.

Other recent studies that should be noted are Felmlee (1984) who useg the NLS panel of Young Women (1968-1973) and continuous-time modelling to analyze the process by which young women leave employment to become unemployed or to exit the labor force. She separates labor force transitions out of employment due to pregnancy and transitions made for other reasons and among the latter category, further separates voluntary from involuntary exits. Meitzen (1986) uses data from the Employment Opportunities Pilot Programs (EOPP) Employers' Survey to estimate a continuous-time model of male and female quitting behavior. The EOPP data, which were collected March-May 1980 include recently hired women with 2.5 or fewer years of tenure in the firm. Meitzen reports that the possibility of quitting a job declines with tenure for men and increases with tenure for women. However, his model includes only one duration parameter, forcing a monotonic relationship between tenure and the likelihood of quitting. Blank (1988) uses PSID data to estimate a competing risks model for transitions into and out of both full-time and parttime work. She finds age, number of children, race, and education to be
important determinants of labor force movements. This research illustrates the importance of treating female labor force decisions within the context of a dynamic model. A single cross-section of data can yield little insight into the heterogeneity among women in their life-cycle patterns of work.

## 1. Methodology

Our theoretical model follows work by others on dynamic labor supply, especially that of Heckman and MaCurdy (1980) and Flinn and Heckman (1982, 1983). 7 These models generalize one-period models of labor force participation in a straightforward manner. Suppose that we define $W_{r}(t)$ as the marginal rate of substitution between goods and leisure if the woman is not working and $W$ ( $t$ ) as her market wage rate. Let $I(t)$ be an index function defined by: ${ }^{8}$
$I(t)=W$
(t) $-W_{r}$
(t).

If I ( $t$ ) $\geq 0$ then the woman chooses to work in the market and we set a work dummy $d(t)=1$. If $I(t)<0$, she chooses not to work and $d(t)=0$. The object of our model is to explain variation in the number of consecutive years a woman is working. Similarly we would like to explain the variation in consecutive years not working.

Given the dynamic nature of female labor supply and the availability of event history information we use continuous time methods to model the determinants of exit rates from work and entry rates into work. Estimation of

[^4]transition rate (or hazard rate) models has become widespread in economics with applications to variety of economic events including unemployment (Lancaster, 1979; Flinn and Heckman, 1983), marriage and divorce (Anderson et al. 1987; Menken et al. 1981, fertility (Newman and McCulloch 1984; Heckman, Hotz, and Walker, 1985) and welface receipt (O'Neill et ale, 1987). Our model focuses on the exit (entrance) probability at year $t$, conditional on having been in (out of) the labor force for $t$ yearg. The NLs data include both completed apells, for which the time in the labor force is measured, and censored spells, for which the end of the labor force spell is not yet observed. Maximum likelihood estimation makes full use of the information both for completed and censored work and non-work spells. We use information on all labor force spells for all women, which implies multiple spells for some women.

If we define the instantaneous unconditional probability of exiting the labor force at time $t$ to be $f_{w}(t)$, where $F_{w}(t)$ is the corresponding cumulative density and $1-F_{W}(t)$ the unconditional survivor probability, then the conditional probability of exiting the labor force at time $t$ given participation up to that time -- called the hazard rate -- is given by:

$$
h_{W}\left(t_{W}\right)=f_{w}\left(t_{W}\right) /\left(1-F_{W}\left(t_{W}\right)\right)
$$

The transition rate from non-work to work $h_{m w}$ can be defined in an analogous manner.

The expected duration of a work spell, e.g., is then:
$\int_{0}^{\infty} \exp \left(-\int_{0}^{t} h_{w}\left(u_{w}\right) d u\right) d t$

In order to study the determinants of the hazard rates across individual
women we have begun with a general functional form of the multiple spell model. (See Yi, et al. (1987), Flinn and Heckman (1982, 1983), Heckman and Singer (1986), and also Alison (1984).) The general form of the conditional hazard for the transition out of work state $j(=w, n w)$ for individual $i$ can be written as:

where the 2 are explanatory variables that include both fixed and time-varying variables and $\theta$ measures individual unobserved heterogeneity. Our model has gpecified the term capturing the effect of duration to be quadratic, with $\lambda_{j 1}$ $=1$ and $\lambda_{j 2}=2$. We initially specified a non-parametric distribution to allow for the effects of unobserved heterogeneity, however, the estimated results differed little from those with no heterogeneity corrections, therefore we present results for models with no heterogeneity controls.

## 2. Degcription of the Data for the Muitiple Spell Model

We analyze the extent to which the labor force behavior of recent cohorts (and its determinants) has actually changed by comparing the early labor force experience of women who were between the ages of 15 and 19 in 1968 with those who were between the ages of 15 and 19 in 1979. We also consider the cohorts ages 20 to 24 in 1968 and 20 to 24 in 1973. Eight-year samples from the National Longitudinal Surveys of Young Women and Youth form the data base for this analysis. For comparability, the 15 to 19 year old NLs Young Women are followed from 1968 through 1975, while the NLS Youth include observations from the period 1979 through 1986. The 20 to 24 year old groups
from the NLS Young Women are followed from 1968 through 1975 and from 1973 through 1980.

Our objective is to explain variation in the number of consecutive years a woman is working (that is, the duration of a work spell) as well as consecutive years not.working. Given the dynamic nature of female labor supply and the availability of event history information we use continuous time methods to model the determinants of exits from work spells as well as from non-work spells.

The NLS data include both completed spells, the duration of which is known, and censored spells, for which the end of the labor force apell is not yet observed. Maximum likelihood estimation makes full use of the information both for completed and censored work and non-work spells. We use information on all spells for all women, which implies multiple spells for some women. 9 The analysis includes thoge persong who eventually leave the gample if we observe at least three years data for them; their work and non-work spelis are treated as being right-censored.

Spell lengths are measured in years; if an individual reports herself to be employed for at least six months of a given year, that year is counted as a work apell year. Otherwise the year is defined as a non-work apell year. ${ }^{10}$

[^5]The explanatory variables, described in Table 17, include both those which are fixed for a given spell and those which vary with time as the spell progresses. Fixed variables consist of years of experience, ${ }^{11}$ age, and years of schooling, each measured at the beginning of the spell. Time-varying variables include a dummy variable equal to one if the woman is enrolled in school during the spell year, two time-varying residential dummies which capture Southern and SMSA residence, and vectors of dichotomous marital status and fertility variables, included to capture changes in household composition. These variables are married to not married, not married to married, and married to married. Women who remain unmarried in a spell year define the reference category for marital status. The fertility variables include firat birth; subsequent birth; no births, but children younger than six present; and, for the women 20 to 24 , the number of children older than six. ${ }^{12}$ Women who have no children are the reference fertility group. The model includeg also spell duration and its square. ${ }^{13}$

[^6]VARIABLE DEFINITIONS FOR WORK AND NON-WORK SPELL MODEL

## Dependent Variables

Work spells: Consecutive years in which respondent worked at least six months
Non-work Consecutive years in which respondent did not spells: work at least six months

## Explanatory Variables

Fixed at beginning of spell:
EXPERIENCE equals cumulative years of work experience
AGE
GRADE highest level of schooling completed
Time-varying:
ENROLLED equals one if enrolled in school during spell year
SOUTH equals one if residing in Southern state CITY CENTER equals one if in SMSA central city MAR TO NOT MAR equals one if marital dissolution occurs during spell year
NOT MAR TO MAR equals one if respondent marries during spell year
MAR TO MAR equals on if respondent remains married during spell year
BIRTH:NO KIDS equals one if respondent has a first birth during a spell year
BIRTH:KIDS equals one if respondent has a second
or higher order birth during a spell year NO BIRTH:KIDS LT6 equals one if respondent has children under age six, but no births during a spell year
ALL KIDS GTG equals the number of children older than six if all children are over 6 in a spell year
SCHOOL CHILD equals one if youngest child reaches age 6 in a spell year

## 3. Empirical Analysis

Given that labor force behavior may differ by race, models are estimated separately for non-black and black women. ${ }^{14}$ Table 18 describes the dependent variables by including the number and mean duration of censored and completed work and non-work spells. Comparing the 15 to. 19 year olds, the most striking difference in spell characteristics is that the mean duration of censored work spells increases by about one year, rising from 3.9 to 4.9 years (of a potential eight) for non-black women and from 3.3 to 4.1 years for black women. For 15 to 19 year olds, the length of completed work spelis is also higher by one-half year for the more recent group. Censored non-work spells have lengthened for the more recent group of women, with mean duration rising about 6 months, except for Non-black 20 to 24 year olds, for whom they have shortened. The 20 to 24 year olds experience longer censored work spells than does the younger age group, and these spell lengths also increased among the more recent cohort.

Tables 19 and 20 provide somewhat more detail regarding the distribution of spells by their duration, first for censored, then for completed spells. Turning to Table 19, the proportion of all censored work spells that last the entire eight-year period has increased for all age and race groupg. For 15 to 19 year old non-black women, this proportion more than doubles from 9.8 percent to 24.0 percent. The proportion of censored non-work spells that are of eight-years duration also increases modestly over time for each age and

[^7]Table 18
Mean Duration and Number of Censored and Completed Spells, by Age Cohort, Race, and Time Period

| Age and Race | Work Spells |  |  |  | Non-Work Spells |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Censored |  | Completed |  | Censored |  | Completed |  |
|  | 1968-75 | 1979-86 | 1968-75 | 1979-86 | 1968-75 | 1979-86 | 1968-75 | 1979-86 |
| 15-19 to 23-27 |  |  |  |  |  |  |  |  |
| Non-Black | $\begin{array}{r} 3.89 \\ 591 \end{array}$ | $\begin{aligned} & 4.94 \\ & 1582 \end{aligned}$ | $\begin{array}{r} 1.81 \\ 636 \end{array}$ | $\begin{aligned} & 2.29 \\ & 1273 \end{aligned}$ | $\begin{array}{r} 3.58 \\ 357 \end{array}$ | $\begin{array}{r} 4.18 \\ 696 \end{array}$ | 1.99 886 | 2.14 1866 |
| Black | $\begin{array}{r} 3.27 \\ 230 \end{array}$ | $\begin{array}{r} 4.08 \\ 563 \end{array}$ | $\begin{array}{r} 1.42 \\ 251 \end{array}$ | $\begin{array}{r} 1.92 \\ 480 \end{array}$ | $\begin{array}{r} 4.44 \\ 165 \end{array}$ | 5.15 356 | 2.23 389 | 2.60 823 |
|  | 1968-75 | 1973-80 | 1968-75 | 1973-80 | 1968-75 | 1973-80 | 1968-75 | 1973-80 |
| 20-24 to 28-32 |  |  |  |  |  |  |  |  |
| Non-Black | $\begin{array}{r} 4.85 \\ 676 \end{array}$ | $\begin{array}{r} 5.35 \\ 569 \end{array}$ | $\begin{array}{r} 2.16 \\ 846 \end{array}$ | $\begin{array}{r} 2.25 \\ 634 \end{array}$ | $\begin{array}{r} 4.72 \\ 493 \end{array}$ | 4.37 329 | 2.02 900 | 1.86 |
| Black | $\begin{array}{r} 4.41 \\ 244 \end{array}$ | $\begin{array}{r} 5.04 \\ 254 \end{array}$ | $\begin{array}{r} 1.90 \\ 301 \end{array}$ | $\begin{array}{r} 1.87 \\ 245 \end{array}$ | $\begin{array}{r} 4.54 \\ 147 \end{array}$ | 4.95 121 | 2.10 362 | 1.99 288 |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

Humber and Proportion of Censored Work and Hon-Hork Spells, by Spell Length

|  | WORK SPELLLS |  |  |  |  |  |  |  | WOW-WORK SPELLS |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 15-19 T0 23-27 |  |  |  | 20-24 10 28-32 |  |  |  | 15-19 10 23-27 |  |  |  | 20-24 т0 28-32 |  |  |  |
|  | Hon-Black |  | Black |  | Non-Black |  | Black |  | Hon-Black |  | Black |  | Mon-Black |  | Black |  |
|  | 1968-75 | 1979-86 | 1968-75 | 1979-86 | 1968-75 | 1973-80 | 1968-75 | 1973-80 | 1968-75 | 1979-86 | 1968-75 | 1979-86 | 1968-75 | 1973-80 | 1968-75 | 1973-80 |
| 1 Year | 190 | 177 | 72 | 103 | 126 | 75 | 52 | 33 | 109 | 186 | 43 | 59 | 95 | 53 | 43 | 22 |
|  | 0.186 | 0.112 | 0.313 | 0.183 | 0.186 | 0.132 | 0.213 | 0.130 | 0.305 | 0.267 | 0.261 | 0.166 | 0.193 | 0.161 | 0.293 | 0.182 |
| 2 Years | 39 | 180 | 13 | 87 | 28 | 49 | 11 | 25 | 48 | 83 | 12 | 32 | 57 | 58 | 14 | 8 |
|  | 0.066 | 0.116 | 0.057 | 0.155 | 0.041 | 0.086 | 0.045 | 0.098 | 0.134 | 0.119 | 0.073 | 0.090 | 0.416 | 0.176 | 0.095 | 0.066 |
| 3 Years | 114 | 173 | 54 | 85 | 86 | 56 | 44 | 27 | 34 | 67 | 9 | 37 | 35 | $36 \cdot$ | - 10 | 12 |
|  | 0.193 | 0.109 | 0.235 | 0.151 | 0.127 | 0.098 | 0.180 | 0.106 | 0.095 | 0.096 | 0.055 | 0.104 | 0.071 | 0.109 | 0.068 | 0.099 |
| 4 Years | 106 | 170 | 25 | 59 | 70 | 37 | 24 | 25 | 27 | 52 | 9. | 18 | 42 | 35 | 8 | 14 |
|  | 0.179 | 0.107 | 0.109 | 0.105 | 0.104 | 0.065 | 0.098 | 0.098 | 0.076 | 0.075 | 0.055 | 0.051 | 0.085 | 0.106 | 0.054 | 0.116 |
| 5 Years | 76 | 167 | 26 | . 58 | 82 | $\cdots 60$ | - 18 | 27 | 38 | . 49. | 14 | 17 | 49 | 32 | 3 | 13 |
|  | 0.129 | 0.105 | 0.113 | 0.103 | 0.121 | 0.105 | 0.074 | 0.106 | 0.106 | 0.070 | 0.085 | 0.048 | 0.099 | 0.097 | 0.020 | 0.107 |
| 6 Years | 55 | 153 | 18 | 40 | 40 | 28 | 28 | 22 | 14 | 41 | 7 | 12 | 40 | 17 | 11 | 3 |
|  | 0.093 | 0.097 | 0.078 | 0.071 | 0.059 | 0.049 | 0.115 | 0.087 | 0.039 | 0.059 | 0.042 | 0.034 | 0.081 | 0.052 | 0.075 | 0.025 |
| 7 Years | 33 | 185 | 13 | 51 | 46 | 26 | 20 | 11 | 13 | 29 | 7 | 9 | 38 | 25 | 11 | 6 |
|  | 0.056 | 0.117 | 0.057 | 0.091 | 0.068 | 0.046 | 0.082 | 0.043 | 0.036 | 0.042 | 0.042 | 0.025 | 0.077 | 0.076 | 0.075 | 0.050 |
| 8 Years | 58 | 380 | 9 | 80 | 198 | 238 | 47 | 84 | 74 | 179 | 67 | 162 | 137 | 73 | 47 | 43 |
|  | 0.098 | 0.240 | 0.039 | 0.142 | 0.293 | 0.418 | 0.193 | 0.331 | 0.207 | 0.257 | 0.406 | 0.455 | 0.278 | 0.222 | 0.320 | 0.355 |
| Total Number of spella | 591 | 4585 | 230 | 563 | 676 | 569 | 244 | 234 | 357 | 696 | 165 | 356 | 493 | 329 | 147 | 121 |

Source: NLS Surveys of Mature Women, Young Women, and Youth.

Table 20
Number and Proportion of Completed Work and Kon-Work Spelts, by spell Length

|  | WORK SPELLS |  |  |  |  |  |  |  | MON-HORK SPELLS |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 15-19 T0 23-27 |  |  |  | 20-24 10 28-32 |  |  |  | 15-19 то 23-27 |  |  |  | 20-24 10 28-32 |  |  |  |
|  | Non-8lack |  | Black |  | Mon-Black |  | Black |  | Non-Black |  | Black |  | Kon-Black |  | Slack |  |
|  | 1968-75 | 1979-86 | 1968-75 | 1979-86 | 1968-75 | 1973-80 | 1968-75 | 1973-80 | 1968-75 | 1979-86 | 1968-73 | 1979-86 | 1968-75 | 1973-80 | 1968-75 | 1973-80 |
| 1 Year | 411 | 550 | 190 | 240 | 399 | 323 | 172 | 156 | 447 | 883 | 186 | 337 | 487 | 399 | 194 | 178 |
|  | 0.565 | 0.431 | 0.701 | 0.500 | 0.482 | 0.509 | 0.571 | 0.637 | 0.435 | 0.471 | 0.419 | 0.409 | 0.541 | 0.653 | 0.536 | 0.618 |
| 2 Years | 146 | 306 | 50 | 138 | 193 | 90 | 50 | 29 | 221 | 451 | 75 | 150 | 171 | 69 | 69 | 33 |
|  | 0.201 | 0.240 | 0.185 | 0.288 | 0.233 | 0.142 | 0.166 | 0.118 | 0.215 | 0.241 | 0.169 | 0.182 | 0.190 | 0.113 | 0.191 | 0.115 |
| 3 Years | 72 | 167 | 17 | 49 | 105 | 87 | 45 | 24 | 151 | 235 | 55 | 116 | 95 | 52 | 30 | 34 |
|  | 0.099 | 0.131 | 0.063 | 0.102 | 0.127 | 0.137 | 0.150 | 0.098 | 0.147 | 0.125 | 0.124 | 0.141 | 0.106 | 0.085 | 0.083 | 0.118 |
| 4 Years | 54 | 131 | 8 | 25 | 62 | 63 | 16 | 22 | 95 | 129 | 43 | 0.66 -68 | 73 | 49 | . 24 | -12 |
|  | 0.074 | 0.103 | 0.030 | 0.052 | 0.075 | 0.099 | 0.053 | 0.090 | 0.092 | 0.069 | 0.097 | 0.080 | 0.083 | 0.080 | 0.066 | 0.042 |
| 5 Years | 26 | 63 | 5 | 14 | 50 | 29 | 8 | 5 | 82 | 82 | 60 | 67 | 42 | 17 | . 29 | 12 |
|  | 0.033 | 0.049 | 0.018 | 0.029 | 0.060 | 0.046 | 0.027 | 0.020 | 0.080 | 0.044 | 0.135 | 0.081 | 0.047 | 0.028 | 0.080 | 0.042 |
| 6 Years | 17 | 42 | . 1 | 9 | 28 | 30 | 8 | 4 | 26 | 65 | 15 | 52 | 17 | 12 | 13 | 13 |
|  | 0.023 | 0.033 | 0.004 | 0.019 | 0.034 | 0.047 | 0.027 | 0.016 | 0.025 | 0.035 | 0.034 | 0.063 | 0.019 | 0.020 | 0.036 | 0.045 |
| 7 Years | 4 | 33 | 0 | 5 | 9 | 12 | 2 | 5 | 6 | 28 | 10 | 35 | 13 | 13 | . 3 | 6 |
|  | 0.005 | 0.026 | 0.000 | 0.010 | 0.011 | 0.019 | 0.007 | 0.020 | 0.006 | 0.015 | 0.023 | 0.043 | 0.014 | 0.021 | 0.008 | 0.021 |
| total Number of Spells | 728 | 1275 | 271 | 480 | 828 | 634 | 301 | 245 | 1028 | 1873 | 444 | 823 | 900 | 611 | 362 | 288 |

Source: NLS Surveys of Mature Women, Young Monen, and Youth.
race group except the non-black 20 to 24 years olds (for whom the mean duration of such spelis also declined).

Table 20 describes these distributions for completed work and non-work spells. The majority of completed spells last only one or two years. The proportion of completed work spells, that end within one year declines for the more recent cohort of 15 to 19 -year olds, but rises slightly for the more recent group of 20 to 24 year olds.

Tables 21 and 22 present the estimates of the multiple spell hazard rate models for women ages 15 to 19 at the beginning of the sample period, for nonblack and for black women, respectively. ${ }^{15}$ The results reported here do not include corrections for unobserved heterogeneity. However, given the heterogeneity among women in measured characteristics for which we do control, we would anticipate no dramatic change in our estimates. ${ }^{16}$

Considering first the estimated risk of leaving a work apell, the effect of prior experience becomes negative and highly gignificant for the more recent cohort, having an insignificant (but positive) effect for the earliex cohort. Yeara of schooling increases the length of the work spell; its effect changes little for the more recent group. Being enrolled in school increases the likelihood of leaving a work spell. Southern residence is significantly associated with longer work duration among the more recent cohort of women.

[^8]Table 21
Work and Non-Work Spell Hazard Rate Model, Non-Black Women 15 to 19

| $1968-75$ <br> $(N=948)$ | $1979-86$ <br> $(N=2278)$ |
| :---: | :---: |

Estimate Standard Estimate Standard Error Error

Work to Non-Work Transition

INTERCEPT
EXPERIENCE/100
AGE/ 100
GRADE/100
ENROLLED
SOUTH
CITY CENTER
MAR TO NOT MAR
NOT MAR TO MAR
MAR TO MAR
BIRTH:NO KIDS
BIRTH: KIDS
NO BIRTH:KIDS LT6
DURATION
DURATION ${ }^{2}$
-8.3268
8.4791
-0.0516
-15.3758
0.3720
0.0175
-0.0740
-0.9052
0.2513
0.1568
0.3624
0.2834
0.0555
0.7682
-0.3554
1.4674
5.8769
2.6851
1.7290
0.1062
0.0921
0.0985
0.5032
0.1538
0.1389
0.2458
0.4571
0.1680
0.1543
0.0538
$-4.2873$
0.5609
-16.5922
4.9633
$3.9815 \quad 2.3382$
$-14.6974 \quad 2.0032$
$0.3517 \quad 0.0751$
$-0.1145 \quad 0.0623$
$-0.1153 \quad 0.0904$
$0.2950 \quad 0.2104$
$0.1598 \quad 0.1232$
$0.1959 \quad 0.0930$
$0.7427 \quad 0.1514$
$0.2513 \quad 0.2172$
0.5071
0.5119
-0.2019
0.0966
0.0718
0.0206

Non-Work to Work Transition

INTERCEPT
EXPERIENCE/ 100
AGE/ 100
GRADE/100
ENROLLED
SOUTH
CITY CENTER
MAR TO NOT MAR
NOT MAR TO MAR
MAR TO MAR
BIRTH:NO KIDS
BIRTH:KIDS
NO BIRTH:KIDS LT6
DURATION
DURATION ${ }^{2}$
-11.2336
6.1230
9.5546
9.7537
-0.3902
-0.0506
-0.0312
-1.9333
-0.5086
-1.0006
-0.5537
-0.6542
-0.5827
1.5103
-0.4699
-134.8135

| 1.3500 | -5.2180 | 0.5313 |
| ---: | ---: | ---: |
| 4.2746 | 9.1165 | 3.5678 |
| 3.0993 | -2.9470 | 2.3064 |
| 2.5012 | 14.2659 | 1.8995 |
| 0.0900 | 0.3409 | 0.0679 |
| 0.0712 | -0.0253 | 0.0495 |
| 0.0757 | -0.1067 | 0.0720 |
| 0.6141 | -0.2234 | 0.2170 |
| 0.1366 | -0.3170 | 0.1298 |
| 0.1214 | -0.2807 | 0.0788 |
| 0.1828 | -0.5056 | 0.1314 |
| 0.3056 | -0.1536 | 0.1334 |
| 0.1375 | -0.5185 | 0.0782 |
| 0.1297 | 0.9241 | 0.0605 |
| 0.0474 | -0.2717 | 0.0184 |

NEGATIVE LOG LIKELIHOOD

Table 22
Work and Non-Work Spell Hazard Rate Model, Black Women 15 to 19
$1968-75$

$(N=395)$$\cdots$| $1973-80$ |
| :---: |
| $(N=929)$ |

Estimate Standard Estimate Standard Error

Error
Work to Non-Work Transition

INTERCEPT
EXPERIENCE/ 100
AGE/100
GRADE/ 100
ENROLLED
SOUTH
CITY CENTER
MAR TO NOT MAR
NOT MAR TO MAR
MAR TO MAR
BIRTH:NO KIDS
BIRTH: KIDS
NO BIRTH:KIDS LT6
DURATION
DURATION ${ }^{2}$

INTERCEPT
EXPERIENCE/100
AGE/100
GRADE/100
ENROLLED
SOUTH
CITY CENTER
MAR TO NOT MAR
NOT MAR TO MAR
MAR TO MAR
BIRTH:NO KIDS
BIRTH:KIDS
NO BIRTH:KIDS LT6
DURATION
DURATION ${ }^{2}$
NEGATIVE LOG LIKELIHOOD

| -40.2922 | 4.1753 | -7.5730 | 1.1368 |
| ---: | ---: | ---: | ---: |
| -1.9012 | 9.4739 | -12.4535 | 7.4253 |
| -1.8386 | 3.9983 | 3.5270 | 3.0831 |
| -9.9666 | 2.3804 | -22.5093 | 2.6122 |
| 0.2314 | 0.1861 | 0.2288 | 0.1180 |
| -0.0822 | 0.1362 | -0.2486 | 0.0934 |
| -0.0360 | 0.1361 | 0.1086 | 0.0982 |
| -0.1759 | 0.5318 | -0.4685 | 0.4940 |
| 0.2840 | 0.2712 | -0.0451 | 0.2486 |
| -0.2460 | 0.2143 | -0.5832 | 0.2099 |
| 0.4410 | 0.3127 | 0.2411 | 0.2591 |
| 0.0807 | 0.3641 | 0.2203 | 0.3005 |
| -0.1908 | 0.1682 | 0.2968 | 0.1177 |
| 2.2082 | 0.3531 | 0.8635 | 0.1322 |
| -1.2719 | 0.1419 | -0.3625 | 0.0407 |

Non-Work to Work Transition

| -10.1327 | 1.6148 | -4.2102 | 0.6631 |
| ---: | ---: | ---: | ---: |
| 2.9780 | 7.4784 | 10.0282 | 5.7408 |
| 16.2712 | 4.1344 | -1.2155 | 3.5189 |
| 9.8664 | -3.2069 | 28.1962 | 3.1888 |
| 0.2737 | 0.1457 | 0.4745 | 0.0944 |
| 0.0879 | 0.1122 | 0.1712 | 0.0740 |
| 0.0045 | 0.1063 | -0.1036 | 0.0808 |
| -1.0006 | 0.5567 | 0.3123 | 0.3423 |
| -0.0861 | 0.2272 | 0.1825 | 0.2123 |
| -0.2089 | 0.1571 | -0.0009 | 0.1404 |
| -0.0690 | 0.2111 | -0.2239 | 0.1695 |
| -0.8459 | 0.3065 | -0.0520 | 0.1754 |
| 0.0062 | 0.1416 | -0.5802 | 0.0909 |
| 1.1128 | 0.1867 | 0.7181 | 0.0846 |
| -0.3365 | 0.0613 | -0.1647 | 0.0228 |
|  |  |  |  |
| -72.5013 |  | 242.7594 |  |

Among the earlier cohort, only a marital dissolution for white women has a statistically significant effect, reducing the probability of exiting a work spell. Among the NLSY, only remaining married matters (statistically), with opposite effects by race. The relative effect of remaining married is to raige the likelihood of exiting a work spell-for-white women and increases the duration of a work spell for black women. Relative to a childess woman, one who gives birth to her first child during a work spell year is more likely to end the work spell. This effect changes little for white women, but weakens for black women. Bearing a second or higher order child while working has no statistically significant effect on leaving work, yet the presence children under age six increases the probability of ending a work spell. Finally, after roughly one month in a work spell, the effect of duration becomes negative. This supports our intuition that cumulative experience within a spell raises otherwise unmeasured returns to continued work participation and reduces the probability of leaving work. The most notable difference between the two time periods is that, for black women, there is a dramatic increase in the effect of schooling on the duration of their work spells.

Turning to the estimates for non-work spells, residing in the South exerts a statistically significant effect to reduce the length of a non-work spell only for the more recent group of black women, while SMSA is not significant. Being older at the beginning of a non-work spell significantly reduces non-work spells only for the earlier cohort. The risk of leaving a spell of non-work increases with experience and schooling. The strength and significance of these effects increases markediy as we compare across the two time periods, and this is especially true for black women.

The consequence of achool enrollment appears to change over time for
white women. Among the earlier cohort of white women, gchool enrollment lengthens a non-work spell, while the opposite holds for the more recent group and for both groups of black women.

The impact of marital status on the non-work gpells of white women appears diminished over time. Relative to remaining single, separating, getting married and staying married inhibit labor market entry for white women. But only among the earlier cohort of black women does marital gtatus bear any significant effect on non-work spells; the effect of a martial dissolution weakly reduces the likelihood of ending a spell of non-work. For each group childbearing reduces the likelihood of leaving a non-work spell. Among white women, these effects decline somewhat in magnitude for the more recent cohort. For black women, a first or higher order birth during a nonwork spell year has an insignificant effect on the likelihood of leaving the spell. The presence of children under 6 matters only for the NLSY cohort. Tablea 23 and 24 present the results for the 20 to 24 year olds, with both groups from the NLS Young Women. These data include eight-years of observations, with women being 28 to 32 in the last sample year. Comparing the two time periods (which are five years apart), the results for white women remain quite similar. The effect of a change in marital status becomes insignificant for the more recent group; this change holds true for the black women as well. Also, Southern residence reduces the duration of a non-work spell among the more recent group.

Among black women, the effect of experience to increase work spell length doubles for the more recent group and the effect of a first birth becomes insignificant. The effect of experience to reduce the length of non-

Table 23
Work and Non-Work Spell Hazard Rate Model, Non-Black Women 20 to 24

|  | $\begin{array}{r} 1968-75 \\ (\mathrm{~N}=1169) \end{array}$ |  |  | $\begin{aligned} & 1973-80 \\ & (N=898) \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: |
|  | Estimate St |  | Estimate <br> Err | dard |
| Work to Non-Work Transition |  |  |  |  |
| INTERCEPT | -5.0761 | 0.8059 | $9-4.6136$ | 0.7383 |
| EXPERIENCE/100 | 1.4867 | 2.2240 | $0-3.6704$ | 3.7134 |
| AGE/ 100 | 1.3657 | 2.1236 | $6 \quad 4.7135$ | 2.1328 |
| GRADE/100 | -9.1898 | 2.0267 | $7-6.6097$ | 2.6660 |
| ENROLLED | 0.3206 | 0.1553 | 30.3883 | 0.1807 |
| SOUTH | -0.1109 | 0.0783 | $3-0.1037$ | 0.0941 |
| CITY CENTER | 0.0016 | 0.0772 | 20.0560 | 0.0946 |
| MAR TO NOT MAR | -0.5848 | 0.3500 | - -0.0196 | 0.2564 |
| NOT MAR TO MAR | 0.4510 | 0.1370 | $0 \quad 0.0835$ | 0.1840 |
| MAR TO MAR | 0.1938 | 0.0986 | $6 \quad 0.2707$ | 0.1236 |
| BIRTH:NO KIDS | 0.5058 | 0.1925 | 5. 0.6653 | 0.1950 |
| BIRTH:KIDS | 0.1874 | 0.2169 | $9 \quad 0.1449$ | 0.2763 |
| NO BIRTH:KIDS LT6 | 0.0861 | 0.1011 | $1 \quad 0.1767$ | 0.1180 |
| ALL KIDS GT6 | -0.5954 | 0.2937 | $7-0.5478$ | 0.2777 |
| DURATION | $0.6067$ | $0.1016$ | $6 \quad 0.3916$ | 0.1123 |
| DURATION ${ }^{2}$ | -0.2384 | $0.0302$ | -0.1659 | 0.0321 |
| Non-Work to Work Transition |  |  |  |  |
| INTERCEPT | -9.3619 | 0.7644 | $4 \quad-5.4088$ | 0.8066 |
| EXPERIENCE/100 | 4.4442 | 1.7529 | 92.4494 | 3.0362 |
| AGE/100 | 9.3705 | 1.8850 | 01.4885 | 2.0226 |
| GRADE/100 | 10.1085 | 1.7183 | $3 \quad 7.3322$ | 1.9973 |
| ENROLLED | 0.0168 | 0.1363 | $3-0.0417$ | 0.2008 |
| SOUTH | 0.0651 | 0.0686 | $6 \quad 0.2705$ | 0.0826 |
| CITY CENTER | -0.0441 | 0.0713 | $3-0.0315$ | 0.0969 |
| MAR TO NOT MAR | -0.3720 | 0.3137 | $7-0.4660$ | 0.2598 |
| NOT MAR TO MAR | -0.5492 | 0.1747 | $7-0.3492$ | 0.2106 |
| MAR TO MAR | -0.9068 | 0.1066 | $6-0.9010$ | 0.1224 |
| BIRTH: NO KIDS | -0.5725 | 0.1464 | $4-0.9995$ | 0.2065 |
| BIRTH: KIDS | -0.3349 | 0.1508 | $8-0.2108$ | 0.1705 |
| NO BIRTH:KIDS LT6 | -0.5124 | 0.0964 | $4-0.8198$ | 0.1057 |
| ALL KIDS GT6 | -0.7700 | 0.3295 | $5-0.6486$ | 0.2371 |
| DURATION | $1.0734$ | $0.0974$ | $4 \quad 0.8917$ | 0.1138 |
| DURATION ${ }^{2}$ | -0.3443 | 0.0292 | -0.2835 | 0.0338 |
| NEGATIVE LOG LIKELIHOOD | D 294 |  | 266. |  |

Table 24
Work and Non-Work Spell Hazard Rate Model, Black Women 20 to 24

|  | $(\mathrm{N}=391)^{1968-75}$ |  | =375) 1973-80 |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Estimate |  | mate St <br> Error |  |
| Work to Non-Work Transition |  |  |  |  |
| INTERCEPT | -7.9106 | 1.6290 | -5.9315 | 1.3226 |
| EXPERIENCE/100 | -6.6746 | 3.5059 | -12.3340 | 6.8219 |
| AGE/100 | 6.9445 | 3.1014 | 4.1276 | 2.6850 |
| GRADE/100 | -12.2650 | 2.6985 | -10.6435 | 2.8484 |
| ENROLLED | 0.3792 | 0.2859 | -0.2963 | 0.3493 |
| SOUTH | -0.2028 | 0.1429 | 0.0364 | 0.1415 |
| CITY CENTER | -0.1898 | 0.1268 | 0.0927 | 0.1475 |
| MAR TO NOT MAR | -1.1458 | 0.5352 | 0.0638 | 0.3383 |
| NOT MAR TO MAR | -0.5909 | 0.3291 | 0.3032 | 0.2574 |
| MAR TO MAR | -0.1743 | 0.1468 | -0.3384 | 0.1962 |
| BIRTH: NO KIDS | 0.6581 | 0.2910 | 0.1473 | 0.4158 |
| BIRTH: KIDS | 0.3560 | 0.2563 | 0.2252 | 0.3041 |
| NO BIRTH:KIDS LT6 | -0.1120 | 0.1707 | 0.0114 | 0.1614 |
| ALL KIDS GT6 | -0.3061 | 0.2941 | -0.3501 | 0.3086 |
| DURATION | 0.8009 | 0.1863 | 0.3858 | 0.1853 |
| DURATION ${ }^{2}$ | -0.3238 | 0.0580 | -0.2102 | 0.0547 |
| Non-Work to Work Transition |  |  |  |  |
| INTERCEPT | -8.3383 | 1.3457 | -4.7726 | 1.1288 |
| EXPERIENCE/100 | 13.7438 | 2.5975 | 5.0175 | 4.9211 |
| AGE/ 100 | 9.1704 | 3.0380 | 0.7193 | 2.9236 |
| GRADE/100 | 17.7011 | 2.7068 | 16.1815 | 3.8195 |
| ENROLLED | 0.0182 | 0.2335 | 0.2849 | 0.2932 |
| SOUTH | 0.4146 | 0.1298 | 0.3182 | 0.1317 |
| CITY CENTER | -0.1843 | 0.1176 | -0.1218 | 0.1356 |
| MAR TO NOT MAR | -0.8821 | 0.4167 | -0.3746 | 0.3294 |
| NOT MAR TO MAR | 0.1095 | 0.2352 | -0.1311 | 0.2886 |
| MAR TO MAR | -0.3702 | 0.1381 | -0.2681 | 0.1547 |
| BIRTH:NO KIDS | -0.1886 | 0.2710 | -0.0238 | 0.3377 |
| BIRTH:KIDS | -0.1872 | 0.2061 | -0.6477 | 0.2805 |
| NO BIRTH:KIDS LT6 | -0.4291 | 0.1566 | -0.2914 | 0.1436 |
| ALL KIDS GT6 | -0.6170 | 0.3347 | -0.7222 | 0.2840 |
| DURATION | 0.9538 | $0.1548$ | 0.5663 | $0.1652$ |
| DURATION ${ }^{2}$ | -0.2758 | 0.0470 | -0.1831 | 0.0476 |
| NEGATIVE LOG LIKELIHOOD | -10.9549 |  | 85.5035 |  |

work spells declines in magnitude and becomes insignificant. The effect of a second or higher order birth significantiy deters labor market entry among the more recent group of black women.

What do these results inform us about intercohort differences in labor force transitions? Considering the dependent variable alone, the length of censored work spells has risen nearly one year (of a potential eight) for both white and black women. The human capital variables, in particular the level of labor market experience and schooling both increase the duration of work spells and hasten exita from a non-work spell. These effecta appear stronger for the more recent group of 15 to 19 year old women, especially, black women. Yet the intercohort differences in responses to demographic variablea appear mixed. Relative to being single, all marital states generally deter exits from non-work spells; and these effects are weaker among the more recent cohorts of women. Yet, the labor supply effects of childbearing appear not to have changed dramatically over time. ${ }^{17}$

[^9]
## IV. DETERMINANTS OF THE PROPORTION OF YEARS WORKED OVER THE LIFETIME

In this section we summarize the results of an analysis of the proportion of years worked since leaving school, or since age 18 for the two youngest cohorts. The perspective here differs from that of the duration model described in the preceding chapter in that we seek-to examine the relationship between lifetime work experience and lifetime fertility history, marriage patterns, education and other factors. This approach does not incorporate time-dependent variables; and rather than using annual data, employs data which summarize past behavior. Consequently, the data requirements are less stringent than for the duration model and we can perform this analysis for all of the cohorts including the mature women.
A. Empirical Results

We eatimate for all cohorts regression models of the determinants of the proportion of the lifetime worked for a given age group, a: $P_{t}^{a}=\beta{ }_{t}^{a} X_{t}^{a}+u_{t}^{a}$,
where $t$ indicates the survey year. Parameters are allowed to vary over time within an age group and across age groups. We focus on a few key variables which are known to be strongly associated with women's labor force participation. They are as follows:

Years of school completed (measured as of the year the cohort reached the stated ages). Since schooling is positively correlated with the expected wage it is likely to be positively associated with cumulative work participation (Mincer, 1962). The strength of the effect may vary over time and between cohorts due to changes in the return to education (and, hence, in the extent to which wages rise with schooling) and to offsetting income

Marital status is measured by a dichotomous variable which equals one if the woman had never married as of the stated year. In the past a career in the labor market often precluded marriage; consequently women who sought careers were less likely to marry (Goldin, 1990). It has become much more commonplace for married women to work and we expect that the relation between lifetime participation and marital status would weaken from one cohort to the next. Contributing to this effect is the increase over time in the proportion of never-married women who have borne children, a development which has narrowed the difference in home responsibilities by marital status. In fact the availability of welfare to non-married women with children creates particular work disincentives for this group.

Differences in fertility are captured by a series of dichotomous measures of numbers of children ever born -- none, one, two, or three -- each of which is relative to the omitted group of four or more children.

Control variables for age and gouthern residence in the stated year are also included. Southern states have lower wage rates and this relative wage difference (and substitution effect) would likely reduce labor force participation in the South. However, husbands incomes are also lower in the South, and this would be a factor pulling in the opposite direction -- that is, increasing married women's participation. Moreover, welfare payments are substantially lower in the South thereby encouraging relatively higher work participation among non-married women in the South compared to the North, particularly those with low potential earnings.

Complete regression results for the determinants of the proportion of years since leaving school (or since age 18) in which a woman worked at least six months are presented in Appendix A. Table 25 sumarizes the effects of

Table 25
Effects of Schooling and Marital Status on the Proportion of Years Horked Since Leaving School, by Age and Year

schooling and marital atatus on the proportion of years worked, utilizing the coefficients drawn from the full set of regreseions. The table enables a comparison of effecta for different cohorts reaching a given age range in different years and the results are shown separately by race. For example, it shows the effect of schooling on lifetime participation for cohorts reaching ages 35 to 39 in 1967, 1972, and 1983. For white women in the 35-39 age group in 1967 a one-year increase in schooling was associated with a nonsignificant increase of 0.0061 in the proportion of years worked; but in 1983 the effect was 0.0217 and it was statistically significant.

In general, an increase in schooling is associated with an increase in the proportion of years spent working. The strength of this effect increases over time, particularly in the late 1970 and early $1980 s$ and it is gomewhat stronger for blacks. The increase over time occurs within a given age group (across cohorts) and within a given cohort as it ages. The within-cohort effect can be seen by reading down the diagonal -- for example, the cohort ages 35 to 39 in 1967 was 40 to 44 in 1972 , 45 to 49 in 1977 and 50 to 54 in 1982. The change over time seems then to be more strongly related to temporal rather than cohort effects. The increase in the return to schooling in the 1980s provides a possible explanation of the pattern.

The effect of martial atatus on the proportion of years worked differs by cohort and by race. In general, the relation between remaining gingle and lifetime work participation is positive initially but becomes considerably weaker from cohort to cohort. Among white women in the oldest cohort examined -- those ages 40 to 44 in 1967 -- the proportion of years worked is about onefifth higher among never married women compared to married women. This pronounced effect is sharply reduced, however, and becomes insignificant among
the youngest cohort of white women.
Among black women marital status bears no atatiatically significant relationship to lifetime work experience, although the effect of never marrying is generally weakly positive among the oldest cohort and weakly negative more recently. However, for 25 to 29 year-old black women, having never married reduces participation and is statistically significant in 1978 and 1987 (Appendix A). The increased tendency for single women to become mothers has been particularly strong among black woman and this factor likely helps explain the observed pattern.

The set of regression results reported in Appendix A shows that bearing children diminishes lifetime participation and the greater the number of children born the lower the proportion of years worked. The negative effect of fertility on lifetime participation is stronger among white women than among black women. Among both races, once women reach their late forties and fifties when children are likely to be of school age (or older) the effect diminishes somewhat, suggesting that there is an intensification of labor force participation after fertility is completed.

Tables 26 and 27 summarize by cohort for white and black women respectively, the estimated effects of fertility on the proportion of years worked. These effects are calculated from the Tables in Appendix A by taking the difference between the effects of, for example, bearing one child relative to having borne no children. Overall, the most consistent pattern for white women is that the negative effect of increasing parity from three to four or more children strengthens considerably during this period. Relative to white women, childbearing generally has weaker effects on the lifetime work participation of black women, especially when considering the effect of the

Table 26
Estinated Effects of Relative Numbers of Children on Proportion of Years Hork (Nonblack Wonen)

| Age | Since Leaving School |  |  |  | Since Age 18 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1967 | 1972 | 1977 | 1982 | 1978 | 1987 |
|  | 1968 | 19775 | 19978 | 19983 |  |  |
|  |  |  | 1979 | 1984 |  |  |
| 25-29 |  |  |  |  |  |  |
| 0 to 1 |  | -0.0951 | -0.1496 |  | -0.1123 | -0.1340 |
| 1 to 2 |  | -0.1625 | -0.1846 |  | -0.1689 | -0.1579 |
| 2 to 3 |  | -0.0948 | -0.0564 |  | -0.1091 | $-0.1480$ |
| 3 to 4+ |  | -0.0024 | -0.1384 |  | -0.1173 | -0.1203 |
| 30-34 |  |  |  |  |  |  |
| 0 to 1 | -0.1145 |  | -0.1506 | -0.0827 |  |  |
| 1 to 2 | -0.1602 |  | -0.1482 | -0.1512 | $\cdots$ |  |
| 2 to 3 | -0.0938 |  | -0.0908 | -0.1266 |  |  |
| 3 to 4+ | -0.0863 |  | -0.0611 | -0.1115 |  |  |
| 35-39 |  |  |  |  |  |  |
| 0 to 1 | -0.1210 | -0.0796 |  | -0.1174 |  | - $\therefore$ |
| 1 to 2 | -0. 1587 | -0.147at |  | -0.0763 |  |  |
| 2 to 3 | -0.0514 | -0.1185 |  | -0.0722 |  |  |
| 3 to 4+ | -0.0624 | -0.0727 |  | -0.1167 |  |  |
| 40-44 |  |  |  | - |  |  |
| 0 to 1 | -0.1427 | -0.1717 | -0.0727 |  |  |  |
| 1 to 2 | -0.0738 | -0.0960 | -0.1422 |  |  |  |
| 2 to 3 | -0.0690 | -0.0573 | -0.0972 |  |  | $\because$ |
| 3 to 4+ | -0.0434 | -0.0622 | -0.0777 |  |  |  |
| 45-49 |  |  |  |  |  |  |
| 0 to 1 |  | -0.0953 | -0.1703 | -0.0423 | - | $\cdots$ |
| 1 to 2 |  | -0.0787 | -0.0795 | -0.1145 |  |  |
| 2 to 3 |  | -0.0564 | -0.0528 | -0.0921 | . |  |
| 3 to 4+ | - | -0.0522 | -0.0543 | -0.0942 |  |  |
| 50-54 |  |  |  |  |  |  |
| 0 to 1 |  |  | -0.0869 | -0.1331 | , | . $\cdot$ |
| 1 to 2 |  |  | -0.0600 | -0.0907 |  |  |
| 2 to 3 |  |  | -0.0453 | -0.0430 | ; |  |
| 3 to 4+ |  |  | -0.0520 | -0.0336 |  | : $\cdot$ |

*Differences between regression coefficients, from Tables 25-30.

Table 27
Estimated Effects of Relative Numbers of Children on Proportion of Years Hork (Black Wormen)

| Age | Since Leaving School |  |  |  | Since Age 18 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \\ & 1979 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \\ & 1984 \end{aligned}$ | 1978 | 1987 |
| 25-29 |  |  |  |  |  |  |
| 0 to 1 |  | 0.0202 | -0.1278 |  | -0.0638 | -0.0770 |
| 1 to 2 |  | -0.0424 | -0.1477 |  | -0.1113 | -0.0879 |
| 2 to 3 |  | -0.1244 | -0.0861 |  | -0.0707 | -0.1142 |
| 3 to 4+ |  | -0.0511 | -0.0661 |  | -0.1409 | -0.0505 |
| 30-34 |  |  |  |  |  |  |
| 0 to 1 | 0.0480 | -0.1034 |  | -0.0331 |  |  |
| 1 to 2 | -0.1297 | -0.0094 |  | -0.0668 |  |  |
| 2 to 3 | -0.1042 | -0.1406 |  | -0.1340 |  |  |
| 3 to $4+$ | -0.0902 | -0.0595 |  | -0.0444 |  |  |
| 35-39 |  |  |  |  |  |  |
| 0 to 1 | -0.0451 | -0.0145 |  | -0.0883 |  |  |
| 1 to 2 | -0.0369 | -0.0502 |  | 0.0157 |  |  |
| 2 to 3 | -0.1389 | -0.0381 |  | -0.0948 |  |  |
| 3 to 4+ | -0.0467 | -0.1476 |  | -0.0953 |  |  |
| 40-44 |  |  |  |  |  |  |
| 0 to 1 | -0.0567 | -0.0281 | -0.0831 |  |  |  |
| 1 to 2 | -0.0867 | -0.0677 | -0.0200 |  |  |  |
| 2 to 3 | -0.0664 | -0.0929 | -0.0270 |  |  |  |
| 3 to $4^{+}$ | -0.0162 | -0.0806 | -0.1895 |  |  |  |
| 45-49 |  |  |  |  |  |  |
| 0 to 1 |  | -0.0581 | -0.0982 | -0.1167 |  |  |
| 1 to 2 |  | -0.0423 | -0.0417 | 0.0583 |  |  |
| 2 to 3 |  | -0.0875 | -0.0870 | -0.0201 |  |  |
| 3 to 4+ |  | -0.0195 | -0.1021 | -0.1854 |  |  |
| 50-54 |  |  |  |  |  |  |
| 0 to 1 |  |  | -0.0466 | -0.0749 | ' | . |
| 1 to 2 |  |  | -0.0288 | -0.0541 |  |  |
| 2 to 3 |  |  | -0.1055 | -0.0442 |  |  |
| 3 to 4+ |  |  | -0.0321 | -0.1360 |  |  |

*Differences between regression coefficients, from Tables 25-30.
first (and sometimes the second) child. Yet the negative effect on participation of the firgt child borne by black women has incrased over time.

Finally, we note that the effect of southern residence on lifetime participation differs considerably between black and white women. Among white women the effect-is typically negligible and not statistically significant. Among black women, however, those who live in the South typically have worked a larger proportion of years and the effect is statistically significant. Since relative wages in the South versus the North are lower among blacks than among whites the relatively higher labor force participation of black women in the South must be attributable to offsetting income effects. As noted, lower husbands incomes and lower welfare benefits in the South would produce less of a disincentive than in the North. The greater extent of marital dissolution and reliance on welfare among black women is consistent with their greater exposure to the effects of differences in potential welfare benefits. B. Accounting for the Increase in Lifetime Participation

A natural question to ask is the extent to which changes in the number of children borne, in schooling, in marital status and in the other factors examined can explain the upward trend in lifetime work experience. We firgt address this question in Table 28 which summarizes these results by decomposing the intercohort changes in predicted lifetime participation into components measuring the effect of changes due to differences in the mean values of the independent variables and changes due to variations in the estimated structure of the participation model (suppressing the age group superscript):

$$
P_{t}=B_{t} X_{t}+u_{t}, \quad \text { for the earliegt year, }
$$

$$
P_{t+1}=B_{t+1} X_{t+1}+u_{t+1} \text {, for the latest year, then at the }
$$

sample means:
$P_{t+1}-P_{t}=B_{t}\left(X_{t+1}-X_{t}\right)+\left(B_{t+1}-B_{t}\right) \quad X_{t}+\left(B_{t+1}-B_{t}\right)\left(X_{t+1}-X_{t}\right)$

The first term represents the changes due to changes in the sample means, the second term corresponds to variations in the parameters and the final term represents the interaction. The results are presented fixat regarding all variables and second focussing on the changes in fertility (and fertility coefficients) alone.

The increases in participation have been most dramatic among white women younger than 40 and among these groups, much of the overall change appears due to changes in the average levels of the independent variables, with much of effect of changes in the means resulting from starkly lower fertility rates. Among younger black women, changes in the estimated coefficients and implied structure of the participation model actually suggest absolute reductions in participation, although lower fertility (and concomitant increases in participation) offset somewhat these reductions.

To focus on cohort differences in lifetime participation, we pool over time women within the same five-year age group and reestimate the regression models. Tables 29 and 30 display these results. These models constrain the effects of the independent variables to be equal within a five-year age group across time periods. For each age group, the earliest cohort is taken as the reference group. Among white women, with one exception, each of the cohort effects is positive, with cohort dummies implying a five to 15 percent higher

Decomposition of Changes in Lifetime Participation

| Age Group | Overall <br> Change | Due to <br> Means |
| :--- | :--- | :--- | | Due to |
| :---: |
| Coefficients $^{\text {b }}$ Interaction $^{c}$ |


| 25 to 29: 1973-78 Nonblack Total Fertility | 0.1143 | $\begin{gathered} 0.0347 \\ 0.0352 \end{gathered}$ | $\begin{gathered} 0.0596 \\ 0.0711 \end{gathered}$ | $\begin{gathered} 0.0200 \\ 0.0080 \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: |
| Black Total | 0.1134 | 0.0425 | 0.0481 | 0.0228 |
| Fertility |  | 0.0226 | 0.0837 | 0.0071 |
| 30 to 34: 1967-83 |  |  |  |  |
| Nonblack Total | 0.1559 | 0.1392 | 0.0189 | -0.0022 |
| Fertility |  | 0.1116 | 0.0338 | 0.0105 |
| Black Total | 0.0472 | 0.0908 | -0.0810 | 0.0374 |
| Fertility |  | 0.0876 | -0.0268 | -0.0136 |
| 35 to 39: 1967-83 |  |  |  |  |
| Nonblack Total | 0.1165 | 0.0600 | 0.0218 | 0.0347 |
| Fertility |  | 0.0468 | 0.0603 | 0.0094 |
| Black Total | 0.0375 | 0.0547 | -0.0962 | 0.0790 |
| Fertility |  | 0.0446 | -0.0095 | 0.0024 |
| 40 to 44: 1967-77 |  |  |  |  |
| Nonblack Total | 0.0462 | 0.0013 | 0.0408 | 0.0041 |
| Fertility |  | -0.0063 | 0.0566 | -0.0041 |
| Black Total | -0.0406 | -0.0334 | -0.0074 | 0.0002 |
| Fertility |  | -0.0125 | 0.0109 | -0.0390 |
| 45 to 49: 1972-82 |  |  |  |  |
| Nonblack Total | 0.0586 | 0.0050 | 0.0482 | 0.0054 |
| Fertility |  | -0.0084 | 0.0706 | -0.0036 |
| Black Total | 0.0013 | -0.0133 | 0.0143 | 0.0003 |
| Fertility |  | -0.0269 | 0.0261 | 0.0021 |
| 50 to 54: 1977-82 |  |  |  |  |
| Nonblack Total | 0.0272 | 0.0062 | 0.0250 | -0.0040 |
| Fertility |  | -0.0047 | 0.0334 | -0.0015 |
| Black Total | -0.0008 | -0.0195 | 0.0323 | -0.0136 |
| Fertility |  | -0.0266 | 0.0375 | -0.0117 |

a $\quad \beta_{t}\left(X_{t+1}-X_{t}\right)$
$c \quad\left(\beta_{t+1}-\beta_{t}\right) X_{t}$
d $\left(\beta_{t+1}-\beta_{t}\right)\left(X_{t+1}-X_{t}\right)$
Fertility results assume means and ocefficients remain constant for
all variables except numbers of children.

Table 29
Determinants of Proportion of Years Worked Since Leaving School

| VARIABLE | Ages 25 to Parameter ESTIMATE | 29 | Ages 30 to PARAMETER ESTIMATE | 34 T-STAT | Ages 35 to PARAMETER ESTIMATE | 39 T-STAT | Ages 40 to PARAMETER ESTIMATE | 44 T-STAT | Ages 45 to PARAMETER ESTIMATE | 49 T-STAT | Ages 50 to PARAMETER ESTIMATE | 54 T-STAT |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| IMTERCEPT | 0.0183 | 0.1430 | -0.0042 | -0.0350 | 0.2190 | 1.4980 | 0.2038 | 1.2040 | 0.1418 | 0.7180 | 0.3918 | 1.4250 |
| AGE | 0.0030 | 0.6770 | 0.0039 | 1.0770 | -0.0021 | -0.5380 | 0.0000 | 0.0110 | 0.0010 | 0.2540 | -0.0036 | -0.6840 |
| SCHOOL ING | 0.0156 | 5.3690 | 0.0158 | 7.5190 | 0.0144 | 6.3200 | 0.0105 | 4.4720 | 0.0131 | 5.3810 | 0.0156 | 5.1410 |
| SOUTH | -0.0117 | -0.8870 | -0.0077 | -0.7240 | 0.0069 | 0.5910 | 0.0101 | 0.8180 | 0.0165 | 1.2810 | 0.0192 | 1.1870 |
| HEVER MARRIED | 0.0285 | 1.4390 | 0.0333 | 1.5920 | 0.1221 | 4.1940 | 0.1629 | 4.7520 | 0.1750 | 4.7160 | 0.2037 | 4.1180 |
| HO CHILDREM | 0.3964 | 11.8040 | 0.4595 | 22.2370 | 0.4048 | 17.9590 | 0.3640 | 16.0350 | 0.3232 | 13.4940 | 0.2664 | 9.0940 |
| ONE CHILD | 0.2681 | 8.1850 | 0.3455 | 17.5730 | 0.2897 | 14.3040 | 0.2275 | 10.9830 | 0.2141 | 10.0430 | 0.1596 | 6.0570 |
| TWO CHILDREN | 0.0990 | 3.0150 | 0.1905 | 11.1190 | 0.1648 | 10.7530 | 0.1325 | 8.8710 | 0.1295 | 8.2640 | 0.0870 | 4.4500 |
| THREE CHILDREN | 0.0223 | 0.6160 | 0.0845 | 4.6480 | 0.0815 | 5.1670 | 0.0605 | 3.9680 | 0.0658 | 4.1250 | 0.0430 | 2.1270 |
| COHORT 1 |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORT2 |  | - | * |  | 0.0170 | 1.2330 | 0.0176 | 1.3130 3.3180 | 0.0253 | 1.8360 3.6250 | 0.0232 | 1.5570 |
| COHOR <br> COHORT4 <br> COHORT5 | 0.0879 | 7.1280 | $\begin{array}{r} -0.0088 \\ 0.0168 \end{array}$ | $\begin{array}{r} -0.6770 \\ 1.2300 \end{array}$ | 0.0367 | 2.7330 | 0.0462 | 3.310 | 0.034 |  | - | - |
| Adjusted R-Square <br> Sanple size <br> Dependent Mean | $\begin{aligned} & 0.2640 \\ & 0.5914 \end{aligned}$ |  | $\begin{aligned} & 0.3119 \\ & 0.5877 \end{aligned}$ |  | $\begin{aligned} & 0.2704 \\ & 0.2408 \end{aligned}$ |  | $\begin{aligned} & 0.2152 \\ & 0.4635 \end{aligned}$ |  |  |  |  |  |
| BLACK HONEN |  |  |  |  |  |  |  |  |  |  |  |  |
|  | Ages 25 to PRRAMETER ESTIMATE |  | Ages 30 to PARAMETER ESTIMATE | 34 Y-STAT | Ages 35 to PARAMETER ESTIMATE | 39 T-STAT | Ages 40 to PARAMETER ESTIMATE | 44 T-STAT | Ages 45 to PARAMETER ESTIMATE | 49 T-STAT | Ages 50 to PARAMETER ESTIMATE | 54 T-STAT |
| VARIABLE |  | T-STAT | ESTIMATE | T-STAT |  | T-STAT |  | T-STAT |  | T-STAT |  | T-STAT |
| INTERCEPT | -0.7791 | -3.3160 | -0.1773 | $\begin{array}{r}0.7800 \\ \hline 1.2620\end{array}$ | 0.1477 | 0.5000 | 0.7404 | 2.2440 | 0.6962 -0.0098 | 1.8650 .12610 | 0.8292 | 1.6610 |
| SCHOOLIHG | 0.0289 | 6.2750 | 0.0292 | 8.0560 | 0.0229 | 5.7560 | -0.0105 | - 2.8550 | -0.0167 | - 4.2500 | -0.0188 | - 4.5160 |
| SOUTH | 0.0308 | 1.3440 | 0.0570 | 2.8970 | 0.0985 | 4.3100 | 0.1208 | 5.2620 | 0.1296 | 5.5430 | 0.1125 | 3.9090 |
| MEVER MARRIED | -0.0594 | -2.2550 | -0.0412 | -1.6180 | -0.0449 | -1.2180 | 0.0272 | 0.5860 | 0.0456 | 0.9290 | 0.0917 | 1.2360 |
| MO CHILDREN | 0.3135 | 7.2950 | 0.3089 | 8.3330 | 0.2939 | 6.7520 | 0.2633 | 6.6860 | 0.2529 | 6.6270 | 0.2481 | 5.6770 |
| OKE CHILD | 0.2514 | 6.5100 | 0.2537 | 8.3210 | 0.2116 | 5.4520 | 0.2106 | 5.7950 | 0.1850 | 5.0320 | 0.1941 | 4.5940 |
| TWO CHILDREN | 0.1453 | 3.8550 | 0.2057 | 7.3210 | 0.1843 | 5.8390 | 0.1552 | 4.7850 | 0.1654 | 4.9540 | 0.1614 | 4.0090 |
| THREE CHILDREN | 0.0459 | 1.1540 | 0.0668 | 2.3670 | 0.1044 | 3.3380 | 0.0896 | 2.5260 | 0.0936 | 2.6600 | 0.0843 | 1.9470 |
| COHORT1 <br> COHORT2 | - | - | - |  |  |  |  |  |  |  |  |  |
| $\begin{aligned} & \text { COHORTZ } \\ & \text { COHORTS } \end{aligned}$ | - |  | * |  | -0.0115 | -0.4290 | -0.0188 | -0.7500 -0.3830 | 0.0040 0.0194 | 0.1610 0.6920 | 0.0230 | 0.8740 |
| cohort4 | ${ }^{*}$ | * | -0.0697 | -2.7780 | -0.0437 | -1.6340 | -0.0106 | - | 0.0194 | 0.6920 | - | . |
| COMORTS | 0.0608 | 2.6550 | -0.0857 | -3.3580 |  |  | - | $\cdots$ | - | - | - | $\cdots$ |
| Adjusted R-Square | 0.1919 |  | 0.2115 |  | 0.1678 |  | 0.1351 |  | 0.1636 |  | 0.1722 |  |
| Sample Size | . 8.888 |  | 1010 |  | 0.789 |  | . 760 |  | . 666 |  | 0.445 |  |
| Dependent Mean | 0.5055 |  | 0.5447 |  | 0.5433 |  | 0.5773 |  | 0.5805 |  | 0.5855 |  |

there remains a strong, statistically significant effect of the passage of time. Even with identical characteristics, women from more recent cohorts are spending a higher proportion of their time at work in the market.

Determinants of Proportion of Years Worked Since Age 18

NON BLACK WOMEN

VARIABLE

INTERCEPT
AGE
SCHOOLING
SOUTH
NEVER MARRIED
NO CHIIDREN
ONE CHILD
TWO CHILDREN
THREE CHILDREN
COHORT5
COHORT6

Adjusted R-Square
Sample Size
Dependent Mean

BLACK WOMEN

|  | AGES 25 to 29 |  |
| :--- | ---: | ---: |
|  | PARAMETER |  |
| VARIABLE | ESTIMATE | T-STAT |
|  |  |  |
| INTERCEPT | -0.0716 | -3.8490 |
| AGE | 0.0222 | 3.3890 |
| SCHOOLING | 0.0307 | 7.7180 |
| SOUTH | 0.0506 | 2.9070 |
| NEVER MARRIED | -0.0725 | -4.0000 |
| NO CHILDREN | 0.3617 | 9.2700 |
| ONE CHILD | 0.2802 | 7.5480 |
| TWO CHILDREN | 0.1831 | 4.9020 |
| THREE CHILDREN | 0.0855 | 2.1110 |
| COHORT5 |  |  |
| COHORT6 |  | 0.0130 |
| Adjusted R-Square |  |  |
| Sample Size | 0.4932 |  |
| Dependent Mean |  |  |

## V. IS THE FEMALE EXPERIENCE-HAGE PROFILE GRONING STEEPER?

This report has examined intercohort changes in cumulated years of work experience among women. We have found that the quantity of work expexience has increased from cohort to cohort. Here we inquire whether the quality of women's work experience has increased. We do this by estimating the relation between wages and years of work experience. If women were investing more in work skills either on-the-job or in school (e.g. by taking more courses in subjects with a stronger market pay-off) then we would expect to see an increase in the effect of experience on wages from one cohort of women to the next.

To investigate the intercohort change in the relation between years of work experience and earnings we have estimated a series of log wage equations, by cohort, for both white and black women. The full set of equations is displayed in Appendix B, Tables B-1 through B-6 where each table presents the results for specific cohort/race groups at the point of reaching a given age. We first summarize the results of these cohort specific regressions and then describe the results of regressions using data pooling the different cohorts.

## A. Cohort/Age/Race Specific Regressions

In the regressions shown in Appendix B experience is defined as years worked six months or more since completing school. An alternative specification also includes the number of years in the past five during which the respondent was not working. This variable is included for all cohorts and survey years except for 1967, for which it could not be measured. Additional explanatory variables include age, years of schooling, southern residence, SMSA residence, and three dummy variables measuring fertility: no children, one child, and two children, each measuring effects relative to three or more
children ever born. In brief our results are as follows.

Among white women, years of work experience generally have statistically significant and positive effects. In instances where years of experience is not significant the model specification typically includes years not working during the past five, a-variable that is becomes highly collinear with total experience at younger ages. The effect of work experience is less consigtent among black women.

The estimated effect of schooling is generally statistically significant and positive for all women. Southern residence dampens wages, particularly among black women. However, this effect is amall and often ingignificant for white women. SMSA residents receive significantly higher earnings.

Not surprisingly, the effects of children on wages are quite variable and are often not statistically significant. The main effect of children appears to operate through effects on years of work experience. As shown in Section IV, the number of children ever born is an important correlate of lifetime work experience. Once work experience is held constant, the expected effect of children on wages is ambivalent. Women with children may seek situations with flexible hours and other amenities compatible with their child care responsibilities and to the extent there are employer costs associated with providing these amenities, wages would be commensurately lower. However, women who work despite the presence of children may well be positively selected, or they may work harder becauge they contribute to their family's support and these factors would lead to a positive relationship between children and wages.

The coefficient on years of experience can be compared across cohorts and within an age group as done in Table 31 utilizing the resulta of

Table 31
Estimated Wage Effects of Years of Experience

| Age | Experience Since Leaving School |  |  |  | Since | Age 18 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{aligned} & 1967 \\ & 1968 \end{aligned}$ | $\begin{aligned} & 1972 \\ & 1973 \end{aligned}$ | $\begin{aligned} & 1977 \\ & 1978 \end{aligned}$ | $\begin{aligned} & 1982 \\ & 1983 \end{aligned}$ | 1978 | 1987 |
| Black |  |  |  |  |  |  |
| 25-29 |  | $\begin{array}{r} 0.0082 \\ 1.026 \end{array}$ | $\begin{array}{r} 0.0165 \\ 1.918 \end{array}$ |  | $\begin{array}{r} 0.0202 \\ 2.019 \end{array}$ | $\begin{array}{r} 0.0522 \\ 6.041 \end{array}$ |
| 30-34 | $\begin{array}{r} 0.0005 \\ 0.074 \end{array}$ |  | $\begin{array}{r} 0.0138 \\ 2.049 \end{array}$ | $\begin{array}{r} 0.0351 \\ 5.058 \end{array}$ |  |  |
| 35-39 | $\begin{array}{r} 0.0025 \\ 0.533 \end{array}$ | $\begin{array}{r} -0.0033 \\ -0.451 \end{array}$ |  | $\begin{array}{r} 0.0352 \\ 5.844 \end{array}$ |  |  |
| 40-44 | $\begin{array}{r} 0.0151 \\ 3.188 \end{array}$ | $\begin{array}{r} 0.0007 \\ 0.151 \end{array}$ | $\begin{array}{r} -0.0142 \\ -2.159 \end{array}$ |  |  |  |
| 45-49 |  | $\begin{array}{r} 0.0048 \\ 0.975 \end{array}$ | $\begin{array}{r} 0.0016 \\ 0.346 \end{array}$ | $\begin{array}{r} -0.0032 \\ -0.478 \end{array}$ |  |  |
| 50-54 |  |  | $0.0083$ | $\begin{array}{r} -0.0014 \\ -0.324 \end{array}$ | - | - |
| NonBlack |  |  |  |  |  |  |
| 25-29 |  | $\begin{array}{r} 0.0161 \\ 2.386 \end{array}$ | $\begin{array}{r} 0.0326 \\ 5.088 \end{array}$ |  | $\begin{array}{r} 0.0348 \\ 5.365 \end{array}$ | $\begin{array}{r} 0.0413 \\ 6.788 \end{array}$ |
| 30-34 | $\begin{array}{r} 0.0230 \\ 4.502 \end{array}$ |  | $\begin{array}{r} 0.0230 \\ 4.867 \end{array}$ | $\begin{array}{r} 0.0337 \\ 6.025 \end{array}$ |  |  |
| 35-39 | $\begin{array}{r} 0.0172 \\ 4.782 \end{array}$ | $\begin{array}{r} 0.0184 \\ 4.244 \end{array}$ |  | $\begin{array}{r} 0.0388 \\ 9.101 \end{array}$ |  |  |
| 40-44 | $\begin{array}{r} 0.0119 \\ 3.900 \end{array}$ | $\begin{array}{r} 0.0189 \\ 5.436 \end{array}$ | $\begin{array}{r} 0.0228 \\ 5.108 \end{array}$ |  |  |  |
| 45-49 | $\begin{array}{r} 0.0151 \\ 5.477 \end{array}$ | $\begin{array}{r} 0.0200 \\ 5.466 \end{array}$ | $\begin{array}{r} 0.0244 \\ 5.530 \end{array}$ |  |  |  |
| 50-54 |  |  | $\begin{array}{r} 0.0140 \\ 4.173 \end{array}$ | $\begin{array}{r} 0.0205 \\ 6.022 \end{array}$ |  |  |

Summary of regression coefficients and t-statistics from Appendix B Tables.
regression models that do not include years out of the labor force in the past five. We find that among white women, the returns to experience increase systematically in size and statistical significance over this period. For example, the estimated effect for women ages 35 to 39 increases from 0.017 in 1967 to 0.039 in -1983. . And while these effecta are quite variable for black women, among black women ages 30 to 34 and 35 to 39 in 1983, the egtimated returns to experience are positive, statiaticaliy significant, and comparable in magnitude to those for white women in the same year.

## B. Reqressions Using Pooled Data

The comparison of coefficients shown in Table 31 suggesta that more recent cohorts have been gaining more in terms of higher pay from an additional year of work experience than was true for earlier generations. However, it is possible that the effect observed is attributable to some other unspecified cohort effect. To estimate the intercohort change in the returns to experience more precisely we have estimated a model pooled across cohorts (and hence across survey groups) for specific age groups. These models focus on the first specification used in the Appendix $B$ tables that includes only the cumulative level of experience and does not include years out of the labor force. All wages are expressed in constant dollars, deflated by the U.S. Bureau of Labor Statistics experimental Consumer Price Index (CPI-U-X1). For this pooled model, the cohorts are combined into five-year age groupg. Dummy variables for the cohort and for the cohort interacted with experience are included as explanatory variables. Each cohort is considered relative to the earliest one relevant for that age group.

The results of these pooled models are shown in Table 32. Among white women the cohort dummy itself is negative in each case, and usually becomes

| NON BLACK HOMEN |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Variable | $\begin{aligned} & \text { Ages } 25 \\ & \text { PARAMETER } \\ & \text { ESTIMATE } \end{aligned}$ | $\begin{aligned} & \text { to } 29 \\ & \text { T-STAT } \end{aligned}$ | $\begin{aligned} & \text { Ages } 30 \\ & \text { PARAMETER } \\ & \text { ESTIMATE } \end{aligned}$ | $\begin{aligned} & \text { to } 34 \\ & \text { T-STAT } \end{aligned}$ | Ages 35 PARAMETER ESTIMATE | $\begin{aligned} & \text { to } 39 \\ & \text { T-STAT } \end{aligned}$ | $\begin{aligned} & \text { Ages } 40 \\ & \text { PARRMETER } \\ & \text { ESTIMATE } \end{aligned}$ | $\begin{aligned} & \text { to } 44 \\ & \text { T-STAT } \end{aligned}$ | Ages 45 PARAMETER ESTIMATE | $\begin{aligned} & \text { to } 49 \\ & \text { T-STAT } \end{aligned}$ | $\begin{aligned} & \text { Ages } 50 \\ & \text { PARMAETER } \\ & \text { ESTIMATE } \end{aligned}$ | $\begin{aligned} & \text { to } 54 \\ & \text { T-STAT } \end{aligned}$ |
| INTERCEPT | 0.5009 | 2.4010 | 0.5844 | 2.4170 | 1.2940 | 4.1430 | 1.1257 | 3.6250 | 1.7004 | 4.1450 | 1.5028 | 4900 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| SCHOOLIHG | 0.0591 | 10.8930 | 0.0783 | 17.6410 | 0.0740 | 15.3690 | 0.0685 | 14.5990 | 0.0749 | 14.3110 | 0.0713 | 10.3240 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| SMSA | 0.1022 | 4.5500 | 0.1684 | 7.6520 | 0.1547 | 6.1660 | 0.1133 | 4.4890 | 0.1248 | 4.6910 | 0.1444 | 4.1130 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| ONE CHILD | 0.0773 | 1.7660 | 0.0785 | 2.1990 | 0.0147 | 0.3840 | 0.0326 | 0.8070 | -0.0118 | -0.2790 | 0.0202 | 0.3680 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORT1 |  |  |  |  |  | - | . ${ }^{\text {U }}$ | - | 228 |  | 0.02 |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORT3 | - | - | 0.093 | 13110 | -0.1253 | -1.7410 | -0.2083 | -3.0240 | -0.1154 | -1.5460 |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORT5 | -0.1380 | -3.2940 | -0.0182 | -2.5290 |  |  | - |  | - |  | - |  |
| COHORT1*EXPERIENCE - - |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORI2*EXPERIENCE | * | - | - | - | - | - | 0.0039 | 0.9780 | 0.0041 | 1.0470 | 0.0040 | 0.9920 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORT4*EXPERIENCE | 0.0159 |  | 0.0045 | 0.6950 | 0.0232 | 4.5650 | . | - | . |  |  |  |
| COHORT5*EXPERIENCE 0.0159 2.1520 0.01281 .8420 - |  |  |  |  |  |  |  |  |  |  |  |  |
| Adjusted R-Square Sample size Dependent Mean | 0.1984 |  | 0.3091 |  | 0.3057 |  | 0.2916 |  | 0.2898 |  |  |  |
|  | . 1291 |  | . 1376 |  | . 1119 |  | . 2980 |  | . 4951 |  |  |  |
|  | 1.9366 |  | 1.9963 |  | 1.9618 |  | 1.9265 |  | 1.9465 |  | 1.9601 |  |
| BLACK MONEN |  |  |  |  |  |  |  |  |  |  |  |  |
|  | ${ }^{\text {Ages } 25}$ | to 29 | Ages 30 | to 34 | Ages 35 | to 39 | Ages 40 | to 44 | Ages 45 to |  | Ages 50 | to 54 |
| VARIABLE ESTIMATE t-STAT ESTIMATE T-StAT ESTIMATE t-STAT ESTIMATE t-STAT ESTIMATE T-STAT ESTIMATE t-STAT |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| SCHOOLING | 0.0590 | 8.8360 | 0.0709 | 11.7530 | 0.1016 | 14.0930 | 0.0868 | 13.1010 | 0.0952 | 13.8610 | 0.0804 | 10.0870 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| SHSA | 0.0878 | 2.4380 | 0.1883 | 5.1730 | 0.1769 | 3.9290 | 0.1930 | 4.1970 | 0.1707 | 3.4600 | 0.1853 | 2.9950 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| ONE CHILO | 0.0288 | 0.6670 | -0.0348: | -0.8490 | 0.0396 | 0.6730 | -0.0329 | -0.5380 | -0.0827 | -1.2980 | 0.0810 | 1.0370 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORT1 | . | . | - |  | - |  | . | ${ }_{*}$ | * | * | * | , |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| cohort3 | + | - | ${ }^{*}$ | * | 0.2601 | 2.3550 | 0.4240 | 3.1140 | 0.0474 | 0.3650 | 0.108 | 1.2270 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORTS | -0.0110 | -0.1780 | -0.2103 | -2.2500 | - 23 | 2.250 | * | - | - | - | * | . |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORT2*EXPERIENCE | - | - | - | - | * | ${ }^{*} 160$ | -0.0111 | -1.7490 | -0.0002 | -0.0420 | -0.0079 | -1.2860 |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| COHORT4*EXPERIENCE | ${ }_{0}^{*} 0069$ | 0.6080 | 0.0181 | 2.0460 | 0.0282 | 3.5320 | - | - | - | - | - | - |
| COHORT5*EXPERIENCE 0.0061 0.6080 0.0333 3.7380 - |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |
| Sample size | $\text { U. } 509$ |  | $568$ |  | $437$ |  | $\begin{aligned} & 874 \\ & 397 \end{aligned}$ |  | $0.4995$ |  | $\begin{array}{r} 0.4444 \\ 225 \end{array}$ |  |
| Dependent Mean | 1.7925 |  | 1.8238 |  | 1.7045 |  | 1.6764 |  | 1.7647 |  | 1.7612 |  |

Note: The years in which each cohort reached ages 25 to 29 is as follows:
A star indicates the cohort which is the reference group in the specific regression.
more negative among higher order (i.e., more recent) cohorts. This implies that, all else equal, more recent cohorts of women gtart at lower wage levels, although not all of these cohort effecta are statistically significant. Moreover, the effect of cohort interacted with experience is positive, and grows larger, indicating that women in more recent cohorts earn higher returns to investments in human capital than their earlier counterparts. For example, for white women 25 to 29 , the effect of experience on earnings increases from 0.0153 for cohort 4 in 1978 by 0.0159 to 0.0302 for cohort 5 in 1983. The cohort differentials in the estimated return to experience are larger and more likely to be statistically significant when the cohorts are farther apart in time. Results based on experience defined since age 18, displayed in Table 33 accord with this findinge.

Among black women younger than 40 , these effects (when significant) imply similar gains in returns to experience as among white women. Yet, for black women 40 to 44 , ftarting wages (all else equal) are higher for more recent cohorts, with returns to experience estimated at lower levels. Although this pattern holds for black women 35 to 39 in 1972 compared with those in 1967 , by 1983 , the return to experience differential becomes positive.

These results provide evidence that work related investments have increased from cohort to cohort among white women, although not necessarily for all cohorts of black women. The negative cohort effect among labor force entrants may be explained by human capital theory which suggesta that on-thejob training would be funded in part through lower wages paid to the trainee (Becker, 1962). Since training is more likely to occur when workers begin their careers, an increase in on-the-job training from one generation to the

Determinants of Log Wages Years of Experience Since Age 18

NON BLACK WOMEN

VARIABLE
INTERCEPT
AGE
EXPERIENCE
SCHOOLING
SOUTH
SMSA
NO CHILDREN
ONE CHILD
I'WO CHILDREN
COHORT5
COHORT6
COHORT5*EXPERIENCE
COHORT6*EXPERIENCE

Ages 25 to 29 PARAMETER ESTIMATE T-STAT
$0.4701 \quad 2.2860$
0.00750 .9860
$0.0339 \quad 4.8570$
0.061713 .4790
$-0.0343-1.7650$
$0.1350 \quad 6.2620$
$0.1413 \quad 2.8950$
0.09691 .9820
0.06671 .3600

*     * 

$-0.0662-1.2150$
0.01011 .2520

Adjusted R-Square
0.2239

Sample Size
Dependent Mean

1953
1.9165

BLACK WOMEN

VARIABLE
INTERCEPT
AGE
EXPERIENCE
SCHOOLING
SOUTH
SMSA
NO CHILDREN
ONE CHILD
TWO CHILDREN
COHORT5
COHORT6
COHORT5*EXPERIENCE
COHORT6*EXPERIENCE

Ages 25 to 29
PARAMETER
ESTIMATE T-STAT
$0.8008 \quad 2.4930$
0.00570 .4820
$0.0200 \quad 1.9510$
$0.0648 \quad 9.2740$
$-0.1343-4.4520$
$0.0968 \quad 2.5840$
$-0.0523-1.0220$
$-0.0575-1.1770$
-0.0785-1.5660
*
*
-0.2601 -3.5820
*
*
$0.0344 \quad 2.8690$

Adjusted R-Square
0.2542

Sample Size
Dependent Mean

720
1.7941
*Designates the cohort that is the reference group for the regression.
next would be expected to generate a greater reduction in wages for new entrants among more recent cohorts. The subsequent steeper rise in earnings with experience reflects the return to the greater level of investment in work skills.

Of course, we cannot determine from this analysis the extent to which women or their employers are responsible for the increase in investment levels: For example, women may have initiated the change by applying for entry into management training programs or through choice of occupations where training is more commonplace. Employers, however, may have become more willing to invest in the training of women either because they have become less prejudiced, or because they are responding to the reduction in turnover and increased labor force attachment exhibited by the average woman, a factor that would have reduced the risk of training women. Since there is so much feedback between the labor force choices of women and the behavior of employers it is probable that both sides contributed to the increase in earnings effects of experience that we have estimated for more recent cohorts of women. The old pattern of flat age-earnings profiles for women -- the dead-end job syndrome -- finally appears to have been overcome, which bodes well for future narrowing of the gender gap in earnings.

Allison, Paul D. "Event History Analysis: Regreasion for Longitudinal Event Data." Sage University Paper Series on Quantitative Applications in the Sócial Sciencea, 07-046. Beveriy Filis: Sage Publicacions, 1984.

Anderson, K.H., Hill, M.A., and Butler, J.S., "Age at Marriage in Malaysia; A Hazard Model of Marriage Timing," Journal of Development Economics 26 (1987): 223-234.

Barnes, William $F$. and Jones, Ethel B., "Difference in Male and Female Quitting," Jourinal of Human Resources 9 (Fall 1974):439-451.

Ben-Porath, Yoram, "Labor Force Participation Rates and the Supply of Labor," Journal of Political Economy, 81 (May-June 1973):697-704.

Blank, Rebecca, "The Role of Part-Time Work in Women's Labor Market Choices Over Time," American Economic Review Papers and Proceedings, May 1989

Cain, Glen, Married Women in the Labor Force. Chicago: University of Chicago Press, 1966.

Center for Human Resource Research. NLS Handbook 1987: The National Longitudinal Surveys of Labor Market Experience. Columbus, Ohio: Center for Human Resource Research, 1987.

Cogan, J. "Married Women's Labor Supply: A Comparison of Alternative Eatimation Procedures," in J. P. Smith ed. Female Labor Supply: Theory and Estimation. Princeton, N.J.: Princeton University Press, 1980.

Corcoran, Mary and Duncan, Greg, "Work History, Labor Force Attachment, and Earnings Differences Between the Races and Sexes," Journal of Human Resources 14 (Winter 1979):3-20.

Donohue, John J., "The Changing Relative Hazard Rates of Young Male and Female Workers," unpublished paper, Northwestern University School of Law, American Bar Foundation, September, 1987.

Ehrenberg, R. and R. Oaxaca, 1976, "Unemployment Insurance, Duration of Unemployment, and Subsequent Wage Gain," American Economic Review 66:734-766.

Even, William E., "Career Interxuptions Following Childbirth," Journal of Itabor Economice, 5 (April 1987):255-277.

Ferber, Marianne A. and Spaeth, Joe L., "Work Characteristics and the MaleFemale Earnings Gap," American Economic Review 74 (May 1984):260-64.

Felmlee, Diane H., "A Dynamic Analysis of Women's Employment Exits," Demography 21 (May 1984):171-183.

Flinn, C. and J. Heckman, "Models for the Analysis of Labor Force Dynamics," Advances in Econometrica (1982) 1:35-95.

Flinn, C. J. and J. J. Heckman, "New Methods for Analyzing Structural Models of Labor Force Dynamics," Journal of Econometrics (1982) 18:115-168
, "Erratum and Addendenum to Volume One: Models for the Analysis of Labor Force Dynamics," Advances_in Econometrics (1983) 2:219223.
"The Likelihood Function for the Multistate Multiepisode model in Molds for the Analysis of Labor Force Dynamics," Advances in Econometrics (1983) 2:225-231.

Goldin, Claudia, "Life Cycle Labor Force Participation of Married Women: Historical Evidence and Implications," Journal of Labor of Economics (1989) 7(1): 20-47 American Women. New York: Oxford University Press, 1990. , Understanding the Gender Gap: An Economic History of

Heckman, J. J. "Shadow Prices, Market Wages and Labor Supply, : Econometrica (1974):679-694.

Heckman, J. J. and T. MaCurdy "A Life Cycle Model of Female Labor Supply," Review of Economic Studies (1980) 47:47-74.

Heckman, J. J. and B. Singer, "Econometric Analysis of Longitudinal," in $Z$. Grilliches and M. D. Intrilligator 1 ed. 7, Handbook of Econometrics, Volume III. 1690-1763. Amsterdam: Elsevier Science Publishers BV, 1986.

Heckman, J. J. and Willis, R. J., "A Beta-Logistic Model for the Analysis of Sequential Labor Force Participation by Married Women." Journal of Political Economy 85 (February 1977): 27-58.

Hill, M. Anne, "Intercohort Differences in Women's Labor Market Transitions," American Economic Review (1990) 80(2):289-292.

Hill, M. Anne and June E. O'Neill, "A Dynamic Model of Women's Work," presented at the Population Association of America's annual meetings in Baltimore, Md., March 1989.

Kiefer, Nicholas M., "Economic Duration Data and Hazard Functions," Journal of Economic Literature (1988) 26(2):646-679.

Killingsworth, Mark R. Labor Supply. Cambridqe: Cambridge Univergity Presg, 1983.
and J. J. Heckman, "Female Labor Supply," in Orley Ashenfelter and Richard Layard, eds. Handbook of Labor Economicas, Volume 1 , 103-204. Amsterdam: North-Holland, 1986.

MaCurdy, Thomas E., "Interpreting Empirical Models of Labor Supply in an Intertemporal Framework with Uncertainty," in James J. Heckman and Burton

Singer, eds., Longitudinal Analygie of Labor Market Data, 111-155. Cambridge: Cambridge Univeraity Press, 1985.

Mincer, Jacob and Solomon Polachek, "Family Investments in Human Capital," Journal of Political Economy 82 (March/April 1974): s76-S108.
(Winter 1978): 118-33.
Mincer, J., and ofek, H. "Interrupted Work Careers: Depreciation and Restoration of Human Capital." Journal of Human Resourcea, (1982) 17(1):3-24.

Nakamura, Alice and Nakamura, Masao, "Dynamic Models of the Labor Force Behavior of Married Women Which Can Be Estimated Using Limited Amounts of Past Information," Journal of Econometrice 27, (1985):273-298.

Newman, J.L. and C.C. McCulloch, "A Hazard Rate Approach to the Timing of Births," Econometrica 52 (1984):939-962.

O'Neill, June, "The Determinants and Wage Effects of Occupational Segregation." Working Paper. Washington, D.C.: Urban Institute, 1983.

Olsen, Randall J. and George Farkas, " Endogenous Covariates in Duration Models and the Effect of Adolescent Childbirth on Schooling," Journal of Human Resources 24(1) (Winter 1989):39-53.

O'Neill, June, "A Time Series Analysis of Women's Labor Force Participation," American Economic_Review 71 (May 1981):76-80.

O'Neill, June, "The Trend in the Male-Female Wage Gap in the United Stater," Journal of Labor Economice 3 (January 1985 Supplement): S91-si16.

O'Neill, June, Douglas A. Wolf and Laurie J. Bassi, "The Duration of Welfare Spells," The Review of Economics and Statistics, Vol. LxIX, No. 2, (May 1987):241-248.

Owen, John D., Workinq Lives. Lexington, MA: Lexington Books, 1986.
Sandell, Steven H. and Shapiro, David, "Work Expectations, Human Capital Accumulation, and the Wages of Young Women," Journal of Human Resources (Summer 1980).

Smith, James P. and Ward, Michael P., "Times-Series Growth in the Female Labor Force, "Journal of Labor Economics 3 (January 1985 Supplement):s59-s90.

Yi, Kei-Mu, Bo Honore, and James Walker, CTM: A Program for the Estimation and Testing of Continuous Time Multi-State Multi-Spell Modelg. Chicago: The University of Chicago, 1987.

APPENDIX A

Determinants of Proportion of Years Worked, Ages 25-29

|  | Since Leaving School |  |  |  | Since Age 18 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| VARIABLE | $1973$ <br> parameter ESTIMATE | T-STAT | 1978 <br> PARAMETER ESTIMATE | T-STAT | 1978 <br> PARAMETER ESTIMATE | T-STAT | $\begin{aligned} & 1987 \\ & \text { PARANETER } \\ & \text { ESTIMATE } \end{aligned}$ | T-STAT |
| INTERCEPT | 0.4020 | 2.0020 | -0.2968 | -1.8190 | -0.0150 | -0.091 | -0.4553 | -3.515 |
| AGE | -0.0030 | -0.4200 | 0.0072 | 1.3160 | 0.0080 | 1.446 | 0.0199 | 4.297 |
| SCHOOL ING | 0.0006 | 0.1380 | 0.0295 | 8.4980 | 0.0014 | 0.405 | 0.0126 | 3.917 |
| SOUTH | -0.0020 | -0.0970 | -0.0230 | -1.4290 | -0.0482 | -2.957 | -0.0104 | -0.769 |
| MEVER MARRIED | 0.0472 | 1.3820 | 0.0049 | 0.2180 | -0.0347 | -1.529 | -0.0664 | -4.228 |
| NO CHILDREN | 0.3547 | 8.0060 | 0.5290 | 9.0330 | 0.5076 | 8.567 | 0.5603 | 13.737 |
| ONE CHILD | 0.2596 | 6.1310 | 0.3793 | 6.5310 | 0.3953 | 6.719 | 0.4263 | 10.568 |
| THO CHILDREN | 0.0971 | 2.2960 | 0.1948 | 3.3580 | 0.2264 | 3.853 | 0.2684 | 6.656 |
| THREE CHILDREN | 0.0024 | 0.0490 | 0.1384 | 2.2550 | 0.1173 | 1.879 | 0.1203 | 2.686 |
| Adjusted R-Square | 0.1537 |  | 0.3675 |  | 0.2227 |  | 0.2688 |  |
| Sample Size | 1204 |  | 1241 |  | 1193 |  | . 1812 |  |
| Dependent Mean | 0.5245 |  | 0.6563 |  | 0.5756 |  | 0.6356 |  |

## BLACK HOMEN

|  | Since Leaving School |  |  |  | Since Age 18 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Variable | $\begin{array}{r} 1973 \\ \text { PARAMETE } \\ \text { ESTIMATE } \end{array}$ | $\mathrm{R}_{\text {t-stat }}$ | 1978 <br> PARAMETER <br> ESIIMATE T-stat |  | $1978$ <br> Parameter estimate | T-STAT | 1987 <br> PARAMETER ESTIMATE | T-Stat |
| Variable |  |  | EStMATE | T-stat |  |  |  |  |
| 1 INTERCEPT | -0.1382 | -0.3740 | -1.3637 | -4.4680 | -1.0417 | -3.699 | -0.5480 | -2.270 |
| AGE | 0.0069 | 0.5270 | 0.0483 | 4.5400 | 0.0388 | 3.945 | 0.0085 | 0.966 |
| SCHOOLIHG | 0.0244 | 3.1310 | 0.0310 | 5.6040 | 0.0184 | 3.618 | 0.0483 | 7.692 |
| SOUTH | -0.0066 | -0.1770 | 0.0534 | 1.8730 | 0.0217 | 0.812 | 0.0639 | 2.790 |
| NEVER MARRIED | -0.0386 | -0.8300 | -0.0617 | -1.9840 | -0.0392 | -1.370 | -0.0925 | -3.961 |
| MO CHILDREN | 0.1977 | 3.1140 | 0.4278 | 7.0450 | 0.3867 | 6.971 | 0.3295 | 5.814 |
| ONE CHILD | 0.2179 | 3.7450 | 0.3000 | 5.5430 | 0.3229 | 6.482 | 0.2526 | 4.549 |
| THO CHILDREN | 0.1755 | 3.2110 | 0.1523 | 2.8200 | 0.2116 | 4.264 | 0.1647 | 2.921 |
| three children | 0.0511 | 0.9210 | 0.0661 | 1.1390 | 0.1409 | 2.627 | 0.0505 | 0.822 |
| Adjusted R-Square | 0.0973 |  | 0.2622 |  |  |  | 0.2458 |  |
| Sample Size | 0395 |  | 492 |  | 458 |  | 653 |  |
| Deperdent Mean | 0.4425 |  | 0.5561 |  | 0.4776 |  | 0.5042 |  |

Table A-2
Determinants of Proportion of Years Horked Since Leaving School, Ages 30 to 34

| NON BLACK MOWEN |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} 196 \\ \text { PARAMET } \end{array}$ |  | $\begin{array}{r} 197 \\ \text { PARAMET } \end{array}$ |  | $\begin{array}{r} 1983 \\ \text { PARAMETE } \end{array}$ |  |
| variable | estimate | t-stat | estimate | t-stat | estimate | T-StAT |
| IMTERCEPT | -0.0538 | -0.2330 | 0.2861 | 1.4570 | -0.2587 | -1.3350 |
| AGE | 0.0055 | 0.7840 | -0.0069 | -1.1660 | 0.0137 | 2.3470 |
| SCHOOLIMG | 0.0148 | 3.5660 | 0.0212 | 5.9570 | 0.0111 | 3.2910 |
| SOUTH | 0.0173 | 0.8140 | -0.0004 | -0.0200 | -0.0318 | -1.8090 |
| hever married | 0.1171 | 2.3650 | 0.0499 | 1.3890 | -0.0178 | -0.6030 |
| MO CHILDREM | 0.4548 | 11.6270 | 0.4507 | 12.1560 | 0.4720 | 11.6440 |
| OWE CHILD | 0.3403 | 8.7650 | 0.3001 | 8.6680 | 0.3893 | 9.8390 |
| TWO CHILDREM | 0.1801 | 6.7870 | 0.1519 | 4.8620 | 0.2381 | 6.3300 |
| THREE CHILOREN | 0.0863 | 3.2720 | 0.0611 | 1.7900 | 0.1115 | 2.7260 |
| Adjusted R-Square | 0.3082 |  | 0.2999 |  | 0.2656 |  |
| Sample size | 747 |  | 1142 |  | 986 |  |
| Dependent Mean | 0.4627 |  | 0.5549 |  | 0.6181 |  |
| 8LACX MOMEH |  |  |  |  |  |  |
|  | 1967 |  | 1978 |  | 1983 |  |
| variable | estimate | t-stat | Estimate | t-stat | Estimate | t-stat |
| INTERCEPT | 0.7216 | 1.4760 | 0.1345 | 0.3790 | -1.1952 | -3.3640 |
| AGE | -0.0106 | -0.7170 | -0.0106 | -0.9770 | 0.0379 | 3.5450 |
| SCHOOL.IMG | -0.0007 | -0.0910 | 0.0496 | 8.3320 | 0.0317 | 5.8890 |
| SOUTH | 0.0644 | 1.4700 | 0.0650 | 2.0410 | 0.0351 | 1.2030 |
| hever married | 0.0438 | 0.6450 | -0.0504 | -1.1910 | -0.0493 | -1.4510 |
| HO CHILDREN | 0.2761 | 2.9370 | 0.3129 | 5.7860 | 0.2783 | 4.6430 |
| ONE CHILO | 0.3241 | 4.8380 | 0.2095 | 4.0790 | 0.2452 | 5.2470 |
| Tho Chtloren | 0.1944 | 2.9860 | 0.2001 | 4.4290 | 0.1785 | 4.0530 |
| three childreh | 0.0902 | 1.5900 | 0.0595 | 1.3690 | 0.0444 | 0.9180 |
| Adjusted R-square | 0.1101 |  |  |  |  |  |
| Sample size | 262 |  | $357$ |  | $389$ |  |
| Dependent Hean | 0.5161 |  | 0.5447 |  | 0.5640 |  |

Determinants of Proportion of Years Worked Since Leaving School, Ages 35 to 39


Table A-4
Determinants of Proportion of Years Worked Since Leaving School, Ages 40 to 44

| MON BLACK MOWEN |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $1967$ <br> parameter |  | $1972$ <br> parameter |  | $\begin{gathered} 1977 \\ \text { PARAMETER } \end{gathered}$ |  |
| variable | estimate | r-Stat | estimate | t-stat | estimate | t-stat |
| INTERCEPT | 0.5212 | 1.8540 | 0.1583 | 0.5310 | -0.1265 | -0.4180 |
| AGE | -0.0064 | -0.9730 | 0.0021 | 0.3060 | 0.0066 | 0.9420 |
| SCHOOLING | 0.0083 | 2.2690 | 0.0087 | 2.0220 | 0.0162 | 3.6330 |
| SOUTH | -0.0075 | -0.3620 | 0.0079 | 0.3570 | 0.0299 | 1.3890 |
| never married | 0.1799 | 2.9710 | 0.1499 | 2.5020 | 0.1385 | 2.4150 |
| ho Chiloren | 0.3289 | 8.9310 | 0.3871 | 9.1540 | 0.3899 | 9.9830 |
| ONE CHILD | 0.1862 | 5.8220 | 0.2154 | 5.7070 | 0.3172 : | 7.9280 |
| TWO CHILDREN | 0.1124 | 4.4690 | 0.1194 | 4.5480 | 0.1750 | 6.7360 |
| THREE CHILDREM | 0.0434 | 1.6310 | 0.0622 . | 2.3580 | 0.0777 | 2.9860 |
| Adjusted R-Square <br> Sample size <br> Dependent Mean | $\begin{array}{r} 0.1711 \\ 859 \\ 0.4442 \end{array}$ |  | $\begin{array}{r} 0.2100 \\ 701 \\ 0.4609 \end{array}$ |  | $\begin{array}{r} 0.2762 \\ 634 \\ 0.4927 \end{array}$ |  |
| BLACK WONEN |  |  |  |  |  |  |
|  | $1967$ <br> PARAMETER |  | $1972$ |  | $1977$ |  |
| - |  |  | PARAMETER |  | ParAmeter |  |
| variable | estimate | T-STAT | estimate | T-STAT | estimate | T-Stat |
| Intercept | 1.1734 | 2.1330 | 0.3903 | 0.6650 | 0.6033 | 1.0260 |
| AGE | -0.0184 | -1.4310 | -0.0034 | -0.2490 | -0.0089 | -0.6470 |
| SCHOOLING | 0.0044 | 0.7630 | 0.0168 | 2.5580 | 0.0133 | 1.7580 |
| SOUTH | 0.1201 | 3.1180 | 0.0916 | 2.2880 | 0.1571 | 3.8240 |
| never married | 0.0926 | 1.2100 | -0.0068 | -0.0620 | -0.0188 | -0.2510 |
| MO CHILDREN | 0.2260 | 4.1650 | 0.2693 | 3.4740 | 0.3196 | 3.2130 |
| ONE CHILD | 0.1693 | 3.2840 | 0.2412 | 3.2400 | 0.2365 | 3.0270 |
| THO CHILDREN | 0.0826 | 1.4730 | 0.1735 | 3.1600 | 0.2165 | 3.7870 |
| three childoren | 0.0162 | 0.2550 | 0.0806 | 1.2850 | 0.1895 | 3.2610 |
| Adjusted R-Square | 0.0921 |  | 0.1281 |  | 0.2042 |  |
| Sample Size | 309 |  | 253 |  | 196 |  |
| Dependent Mean | 0.6017 |  | 0.5482 |  | 0.5609 |  |

Table A-5
Determinants of Proportion of Years Horked Since Leaving School, Ages 45 to 49

|  | $\begin{gathered} 1972 \\ \text { PARAMETER } \end{gathered}$ |  | $1977$ <br> parameter |  | $\begin{gathered} 1982 \\ \text { PARAMETER } \end{gathered}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| variable | Estimate | t-Stat | estimate | T-STAT | estimate | t-stat |
| INTERCEPT | 0.4354 | 1.3500 | 0.2355 | 0.6950 | -0.2705 | -0.7230 |
| AGE | -0.0043 | -0.6410 | 0.0007 | 0.0990 | 0.0083 | 1.0550 |
| SCHOOLING | 0.0112 | 2.9130 | 0.0096 | 2.2430 | 0.0204 | 4.2780 |
| SOUTH | 0.0029 | 0.1380 | 0.0118 | 0.5300 | 0.0340 | 1.4230 |
| hever married | 0.2068 | 3.2230 | 0.1578 | 2.5470 | 0.1346 | 1.9770 |
| NO CHILDREM | 0.2826 | 7.5570 | 0.3569 | 8.3090 | 0.3431 | 7.4840 |
| ONE CHILD | 0.1872 | 5.8620 | 0.1866 | 4.8720 | 0.3008 | 6.8460 |
| TWO CHILDREM | 0.1086 | 4.2440 | 0.1071 | 3.9830 | 0.1863 | 6.3270 |
| three children | 0.0522 | 1.9300 | 0.0543 | 2.0110 | 0.0942 | 3.2420 |
| Adjusted R-Square <br> Semple size <br> Dependent Mean | $\begin{aligned} & 0.1562 \\ & 733 \\ & 0.4573 \end{aligned}$ |  | $\begin{array}{r} 0.1956 \\ 660 \\ 0.4866 \end{array}$ |  | $\begin{array}{r} 0.2465 \\ 537 \\ 0.5164 \end{array}$ |  |

BLACK WOMEN

| Variable | $1972$ <br> parameter |  | $\begin{gathered} 1977 \\ \text { PARAMETER } \end{gathered}$ |  | 1982 <br> PaRAMETER |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | estimate | T-Stat | EStIMATE | r-stat |  | t-stat |
| InTERCEPT | 1.0005 | 1.6150 | 0.4860 | 0.6460 | 0.8464 | 1.1490 |
| AGE | -0.0153 | -1.1820 | -0.0034 | -0.2650 | -0.0146 | -0.9380 |
| SCHOOLING | 0.0131 | 2.1860 | 0.0186 | 2.9790 | 0.0203 | 2.5210 |
| SOUTH | 0.1457 | 3.8390 | 0.0711 | 1.8130 | 0.1938 | 4.1390 |
| mever married | 0.1247 | 1.5160 | 0.0481 | 0.3440 | -0.0026 | -0.0350 |
| MO CHILDREN | 0.2074 | 3.9020 | 0.3290 | 4.3350 | 0.2639 | 2.9850 |
| ONE CHILD | 0.1493 | 2.8760 | 0.2308 | 3.3380 | 0.1472 | 9.6370 |
| THO CHILDREN | 0.1070 | 1.8550 | 0.1891 | 3.5820 | 0.2055 | 3.1790 |
| THREE CHILDREM | 0.0195 | 0.3140 | 0.1021 | 1.7330 | 0.1854 | 2.8570 |
| Adjusted R-Square | 0.1251 |  | 0.1796 |  | 0.2119 |  |
| Sample size | 273 |  | 232 |  | 159 |  |
| Dependent Mean | 0.5870 |  | 0.5674 |  | 0.5884 |  |

Table A-6
Determinants of Proportion of Years Worked Since Leaving School, Ages 50 to 54

| NON BLACK WOMEN |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | $\begin{gathered} 1977 \\ \text { PARAMETER } \end{gathered}$ |  | $\begin{gathered} 1982 \\ \text { PARAMETER } \end{gathered}$ |  |
|  | ESTiMATE | T-sini | ESTIMATE | T-sini |
| IHTERCEPT | 0.5587 | 1.5370 | 0.2169 | 0.5090 |
| AGE | -0.0074 | -1.0760 | 0.0011 | 0.1340 |
| SCHOOLIHS | 0.0182 | 4.6470 | 0.0124 | 2.5460 |
| SOUTH | 0.0208 | 0.9760 | 0.0170 | 0.6780 |
| NEVER MARRIED | 0.2213 | 3.3370 | 0.1745 | 2.3060 |
| MO CHILDREN | 0.2502 | 6.7170 | 0.2904 | 6.0570 |
| ONE CHILD | 0.1633 | 4.9910 | 0.1573 | 3.5020 |
| THO CHILDREN | 0.0973 | 3.7230 | 0.0766 | 2.5730 |
| THREE CHILDREN | 0.0520 | 1.9110 | 0.0336 | 1.1050 |
| Adjusted R-Square | 0.1568 |  | 0.1385 |  |
| Sample Size | 725 |  | 552 |  |
| Dependent Mean | 0.4781 |  | 0.5057 |  |
| BLACK HOHEN |  |  |  |  |
|  | $\begin{gathered} 1977 \\ \text { PARANETER } \end{gathered}$ |  | 1962 <br> parameter |  |
| VARIABLE | ESTIMATE | T-STAT | ESTIMATE | T-STAT |
| IHTERCEPT | 1.2024 | 1.7180 | 0.3512 | 0.4810 |
| AGE | -0.0184 | -1.3890 | -0.0028 | -0.2030 |
| SCHOOLING | 0.0167 | 2.9530 | 0.0220 | 3.5020 |
| SOUTH | 0.1292 | 3.2390 | 0.0894 | 2.1270 |
| NEVER MARRIED | 0.0835 | 0.8750 | 0.1636 | 1.2420 |
| NO CHILDREN | 0.2130 | 3.8500 | 0.3092 | 4.0820 |
| ONE CHILD | 0.1664 | 3.0890 | 0.2343 | 3.2770 |
| THO CHILDREN | 0.1376 | 2.2760 | 0.1802 | 3.3480 |
| THREE CHILDREN | 0.0321 | 0.5140 | 0.1360 | 2.2300 |
| Adjusted R-Square | 0.1329 | : | 0.2208 |  |
| Sample Size | 244 |  | 200 |  |
| Dependent Hean | 0.5859 |  | 0.5850 |  |

APPENDIX B

Table 8-1
Determinants of Log Wages, Ages 25-29

## Years Since Age 18

## NON BLACK HOMEN

| VARIABLE | 1978 (a) |  |  |  | 1978 (b) |  | 1987 (b) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Parameter |  | 1987 (a) PARAMETER |  | parameter |  | PARAMETERESTIMATEt-STAT |  |
|  | estimate | t-stat | ESTIMATE | T-Stat | EStImate | t-stat |  |  |
| IWTERCEPT | 0.6887 | 2.414 | 0.2407 | 0.883 | 0.6648 | 2.339 | 0.1197 | 0.437 |
| AGE | 0.0110 | 1.008 | 0.0075 | 0.732 | 0.0186 | 1.656 | 0.0224 | 2.000 |
| EXPERIENCE | 0.0348 | 5.365 | 0.0413 | 6.788 | 0.0176 | 1.898 | 0.0138 | 1.327 |
| SCHOOLING | 0.0494 | 8.363 | 0.0710 | 10.884 | 0.0470 | 7.902 | 0.0691 | 10.602 |
| SOUTH | -0.0390 | -1.428 | -0.0353 | -1.346 | -0.0429 | -1.575 | -0.0393 | -1.503 |
| SMSA | 0.1324 | 4.587 | 0.1323 | 4.409 | 0.1331 | 4.630 | 0.1342 | 4.488 |
| NO CHILDREN | 0.0221 | 0.330 | 0.2069 | 3.117 | 0.0137 | 0.206 | 0.1775 | 2.659 |
| OFE CHILD | -0.0058 | -0.086 | 0.1569 | 2.379 | -0.0084 | -0.124 | 0.1341 | 2.029 |
| TWO CHILOREN | -0.0389 | -0.561 | 0.1243 | 1.892 | -0.0327 | -0.474 | 0.1147 | 1.750 |
| years out in past five |  |  |  |  | -0.0413 | -2.603 | -0.0555 | -3.245 |
| Adjusted R-Square | 0.2154 |  | 0.2305 |  | 0.2219 |  | 0.2363 |  |
| Semple size | 694 |  | 1256 |  | 694 |  | 1256 |  |
| Dependent Mean | 1.9667 |  | 1.9097 |  | 1.9667 |  | 1.9097 |  |

BLACK HOMEN

| VARIABLE | 1978 (a) |  | 1987 (a) |  | 1978 (b) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | PARAMETER |  | Parameter ESIIMATE | t-stat | Parameter Estimate | T-STAT | $\begin{aligned} & 1987 \text { (b) } \\ & \text { PARAMETER } \end{aligned}$ | T-STAT |
|  |  |  |  |  |  |  |  |  |
| IMTERCEPT | 1.2542 | 2.719 | 0.2528 | 0.591 | 1.2543 | 2.715 | 0.1656 | 0.386 |
| AGE | -0.0033 | -0.195 | 0.0125 | 0.794 | -0.0024 | -0.138 | 0.0245 | 1.454 |
| EXPERIEMCE | 0.0202 | 2.019 | 0.0522 | 6.041 | 0.0173 | 1.167 | 0.0250 | 1.511 |
| SCHOOLIMG | 0.0487 | 5.772 | 0.0781 | 7.112 | 0.0487 | 5.756 | 0.0778 | 7.098 |
| SOUTH | -0.2112 | -4.818 | -0.1019 | -2.503 | -0.2120 | -4.817 | -0.1137 | -2.771 |
| SMSA | 0.0830 | 1.675 | 0.0820 | 1.513 | 0.0837 | 1.684 | 0.0860 | 1.592 |
| NO CHILDREN | 0.0778 | 1.080 | -0.1297 | -1.843 | 0.0758 | 1.044 | -0.1454 | -2.058 |
| OWE CHILD | 0.0313 | 0.481 | -0.1180 | -1.710 | 0.0298 | 0.455 | -0.1284 | -1.861 |
| THO CHILOREK | 0.0289 | 0.424 | -0.1527 | -2.171 | 0.0281 | 0.411 | -0.1576 | -2.245 |
| years out in past five |  |  |  |  | -0.0064 | -0.269 | -0.0493 | -1.930 |
| Adjusted R-Square | 0.2843 |  | 0.2466 |  | 0.2819 |  | 0.2513 |  |
| Sample size | 276 |  | 445 |  | 276 |  | 445 |  |
| Dependent Mean | 1.8626 |  | 1.7757 |  | 1.8626 |  | 1.7757 |  |

Table B-2
Determinants of Log Wages, Ages 25-29

Years Since Leaving school

NOM BLACK WONEM

| Variable | $1973 \text { (a) }$ <br> Parameter |  | $1978 \text { (a) }$ |  | 1973 (b) PARAMETER estimate | T-STAT | $1978$ <br> PARAMETER ESTIMATE | (b) T-STAT |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| INTERCEPT | 0.5502 | 1.654 | 0.4279 | 1.560 | 0.8801 | 2.735 | 0.5111 | 1.863 |
| AGE | 0.0202 | 1.651 | 0.0158 | 1.459 | 0.0215 | 1.837 | 0.0212 | 1.946 |
| EXPERIENCE | 0.0161 | 2.386 | 0.0326 | 5.088 | -0.0055 | -0.774 | 0.0177 | 2.174 |
| SCHOOLING | 0.0515 | 5.478 | 0.0615 | 9.081 | 0.0457 | 5.066 | 0.0538 | 7.453 |
| SOUTH | -0.1537 | -4.525 | -0.0427 | -1.566 | -0.1388 | -4.258 | -0.0447 | -1.646 |
| SHSA | 0.0626 | 1.691 | 0.1299 | 4.495 | 0.0657 | 1.854 | 0.1308 | 4.550 |
| HO CHILDREN | 0.1784 | 2.917 | 0.0472 | 0.713 | 0.0550 | 0.900 | 0.0294 | 0.445 |
| ONE CHILD | 0.1184 | 1.922 | 0.0189 | 0.281 | 0.0556 | 0.934 | 0.0085 | 0.127 |
| TWO CHILDREM | 0.0095 | 0.148 | -0.0362 | -0.523 | -0.0029 | -0.047 | -0.0313 | -0.455 |
| YEARS OUT IN PAST FIVE |  |  |  |  | -0.0974 | -7.070 | -0.0411 | -2.930 |
| Adjusted R-Square | 0.1891 |  | 0.2114 |  | 0.2570 |  | 0.2198 |  |
| Sample size | 545 |  | 714 |  | 545 |  | 714 |  |
| Dependent Mean | 1.9654 |  | 1.9700 |  | 1.9654 |  | 1.9700 |  |

BLACK WOMEN

| VARIABLE | $1973 \text { (a) }$ <br> Parameter |  | $\begin{aligned} & 1978 \\ & \text { PARANETER } \\ & \text { ESTIWATE } \end{aligned}$ | (a) T-STAT | 1973 (b) PARAMETER ESTIMATE | T-STAT | $1978$ <br> PARAMETER estimate | (b) T-STAT |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| IMTERCEPT | 0.6318 | 1.308 | 1.1502 | 2.678 | 1.0120 | 2.131 | 1.2028 | 2.765 |
| AGE | 0.0129 | 0.765 | -0.0013 | -0.079 | 0.0110 | 0.679 | -0.0006 | -0.037 |
| EXPERIENCE | 0.0082 | 1.026 | 0.0165 | 1.918 | -0.0085 | -0.971 | 0.0115 | 1.062 |
| SCHOOLING | 0.0690 | 6.297 | 0.0536 | 6.084 | 0.0577 | 5.285 | 0.0517 | 5.649 |
| SOUTH | -0.3035 | -5.944 | -0.2144 | -5.107 | -0.3102 | -6.307 | -0.2150 | -5.118 |
| SHSA | 0.1153 | 1.952 | 0.0713 | 1.483 | 0.1231 | 2.163 | 0.0723 | 1.500 |
| NO CHILDREH | 0.0200 | 0.305 | 0.1107 | 1.650 | -0.0100 | -0.157 | 0.1010 | 1.478 |
| OWE CHILD | 0.0436 | 0.634 | 0.0560 | 0.913 | 0.0171 | 0.257 | 0.0500 | 0.809 |
| THO CHILDREN | 0.1079 | 1.752 | 0.0639 | 0.997 | 0.0775 | 1.296 | 0.0607 | 0.944 |
| YEARS OUT IN PAST FIVE |  |  |  |  | -0.0737 | -3.974 | -0.0150 | -0.780 |
| Adjusted R-Square | 0.3982 |  | 0.2798 |  | 0.4421 |  | 0.2788 |  |
| Semple size | $197$ |  | 290 |  | 197 |  | $290$ |  |
| Dependent Hean | 1.7814 |  | 1.8589 |  | 1.7814 |  | 1.8589 |  |

Determinants of Log Wages, Ages 30-34

| HON BLACK WOHEN |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| VARIABLE | 1967 (a) <br> PARAMETER <br> ESTIMATE T-STAT |  |  | $\begin{aligned} & 1978 \\ & \text { PARAMETER } \\ & \text { ESTIMATE } \end{aligned}$ | (a) T-STAT | $1983 \text { (a) }$ <br> PARAMETER |  | 1978 (b) <br> PARAMETER |  | $\begin{aligned} & 1983 \text { (b) } \\ & \text { PARAMETER } \end{aligned}$ |  |
|  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  | ESTIMATE |  | T-StAT | ESTIMATE | T-STAT | ESTIMATE | T-STAT |
| INTERCEPT |  | 1.5088 | 3.006 |  | 0.0902 | 0.243 | 0.4898 | 1.272 | 0.3289 | 0.897 | 0.5128 | 1.344 |
| AGE |  | -0.0271 | -1.801 | 0.0184 | 1.559 | -0.0007 | -0.059 | 0.0229 | 1.975 | 0.0096 | 0.778 |
| EXPERIENCE |  | 0.0230 | 4.502 | 0.0230 | 4.867 | 0.0337 | 6.025 | 0.0070 | 1.242 | 0.0168 | 2.273 |
| SCHOOLING |  | 0.0745 | 8.709 | 0.0767 | 10.247 | 0.0788 | 11.100 | 0.0652 | 8.486 | 0.0703 | 9.442 |
| SOUTH |  | -0.1296 | -2.955 | -0.0826 | -2.478 | -0.0009 | -0.026 | -0.0944 | -2.885 | -0.0074 | -0.211 |
| SMSA |  | 0.1808 | 3.841 | 0.1391 | 4.077 | 0.1906 | 5.234 | 0.1401 | 4.192 | 0.1879 | 5.211 |
| NO CHILDREM |  | 0.0396 | 0.640 | 0.0954 | 1.720 | 0.0920 | 1.463 | 0.0474 | 0.859 | 0.0484 | 0.761 |
| ONE CHILD |  | 0.0210 | 0.287 | 0.1253 | 2.251 | 0.0739 | 1.139 | 0.0987 | 1.802 | 0.0395 | 0.608 |
| TWO CHILDREN |  | 0.0446 | 0.785 | 0.0478 | 0.969 | 0.0342 | 0.562 | $\begin{array}{r} 0.0363 \\ -\cap .0601 \end{array}$ | $\begin{array}{r} 0.750 \\ -4.967 \end{array}$ | $\begin{array}{r} 0.0161 \\ -0.0666 \end{array}$ | $\begin{array}{r} 0.265 \\ -3.479 \end{array}$ |
| YEARS OUT IN PAST | FIVE |  |  |  |  |  |  | -0.0691 |  |  |  |
| Adjusted R-Square |  | 0.3833 |  | 0.2849 |  | 0.2967 |  | 0.3142 |  | 0.3103 |  |
| Sample Size |  | . 233 |  | . 564 |  | 573 |  |  |  | $573$ |  |
| Dependent Mean |  | 1.8370 |  | 2.0231 |  | 2.0545 |  | 2.0231 |  | 2.0545 |  |
| BLACK WOMEH |  |  |  |  |  |  |  |  |  |  |  |
|  | - | 1967 (a) |  | 1978 (a) |  | 1983 (a) |  | 1978 (b) |  | 1983 (b) |  |
|  |  | PARAMETER |  | PARAMETER |  | PARAMETER |  | parameter <br> ESTIMAYE T-STAT |  | PARAMETER ESTIMATE |  |
| VARIABLE |  | ESTIMATE | T-STAT | ESTIMATE | T-STAT | ESTIHATE | T-STAT |  |  | T-STAT |
| INTERCEPT |  | 1.2114 | 1.583 | 0.8047 | 1.406 | 1.2342 | 2.267 | 0.9393 | 1.673 |  | 1.1228 | 2.090 |
| AGE |  | -0.0127 | -0.530 | 0.0087 | 0.485 | -0.0210 | -1.223 | 0.0144 | 0.818 | -0.0095 | -0.545 |
| EXPERIENCE |  | 0.0005 | 0.074 | 0.0138 | 2.049 | 0.0351 | 5.058 | -0.0011 | -0.141 | 0.0213 | 2.565 |
| SCHOOL.ING |  | 0.0791 | 5.852 | 0.0617 | 5.872 | 0.0736 | 8.345 | 0.0514 | 4.762 | 0.0663 | 7.343 |
| SOUTH |  | -0.3286 | -4.343 | -0.3096 | -5.763 | -0.1410 | -3.097 | -0.3182 | -6.045 | -0.1348 | -3.005 |
| SHSA |  | 0.1098 | 1.369 | 0.1834 | 3.090 | 0.2237 | 3.898 0.348 | 0.1871 | 3.221 -0.526 | 0.2356 | 4.162 0.105 |
| MO CHILDREN |  | -0.2133 | -1.680 | -0.0022 | -0.034 | 0.0255 | 0.348 | -0.0343 | -0.526 | 0.0076 | 0.105 |
| ONE CHILD |  | 0.0185 | 0.196 | 0.0084 | 0.110 | -0.0626 | -1.064 | -0.0090 | -0.120 | -0.0811 | 1.393 |
| TWO CHILDREN |  | 0.0913 | 0.950 | 0.0032 | 0.052 | 0.0214 | 0.375 | 0:0057 | 0.094 | 0.0160 | 0.286 -2.958 |
| YEARS OUT IN PAST | five |  |  |  |  |  |  | -0.0740 | -3.111 | -0.0780 | -2.958 |
| Adjusted R-Square |  | 0.4615 |  | 0.3724 |  | 0.3739 |  | 0.3991 |  | 0.3935 |  |
| Sample size |  | 113 |  | $\begin{array}{r} 204 \\ 1.8918 \end{array}$ |  | $\begin{array}{r} 249 \\ 1.9200 \end{array}$ |  | $\begin{array}{r} 204 \\ 1.8918 \end{array}$ |  | $\begin{array}{r} 249 \\ 1.9200 \end{array}$ |  |
| Dependent Mean |  | 1.5169 |  |  |  |  |  |  |  |  |  |  |


| NON BLACK HOMEM |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| VARIABLE | 1967 (a) Parameter estimate | T-Stat | 1972 (a) Parameter ESTIMATE | T-STAT | 1983 (a) Parameter estimate | t-stat | 1972 (b) parameter estimate t-stat |  | 1983 (b) parameter estimate t-stat |  |
|  |  |  |  |  |  |  |  |  |  |  |
| IHTERCEPT | 1.1105 | 2.196 | 0.9650 | 1.759 | 1.0867 | 2.246 | 1.0277 | 1.843 | 1.2735 | 2.656 |
| AGE | -0.0083 | -0.625 | -0.0042 | -0.284 | -0.0184 | -1.387 | -0.0044 | -0.299 | -0.0150 | -1.149 |
| EXPERIENCE | 0.0172 | 4.782 | 0.0184 | 4.244 | 0.0388 | 9.101 | 0.0164 | 3.135 | 0.0248 | 4.554 |
| SCHOOLING | 0.0683 | 9.210 | 0.0672 | 7.563 | 0.0793 | 10.158 | 0.0653 | 6.979 | 0.0713 | 8.979 |
| SOUTH | -0.0636 | -1.575 | -0.0760 | -1.746 | -0.1617 | +4.241 | -0.0721 | -1.640 | -0.1556 | -4.135 |
| SMSA | 0.0965 | 2.269 | 0.1421 | 3.088 | 0.1860 | 4.802 | 0.1464 | 3.147 | 0.1929 | 5.046 |
| NO Childrek | -0.0232 | -0.386 | 0.0595 | 0.877 | 0.0637 | 1.096 | 0.0634 | 0.929 | 0.0734 | 1.279 |
| OHE CHILD | -0.0499 | -0.698 | 0.0698 | 0.969 | 0.0247 | 0.423 | 0.0709 | 0.983 | 0.0382 | 0.662 |
| THO Children | -0.0283 | -0.639 | -0.0337 | -0.689 | 0.0272 | 0.591 | -0.0329 | -0.672 | 0.0431 | 0.945 |
| years out in the past five |  |  |  |  |  |  | -0.0114 | -0.656 | -0.0715 | -4.064 |
| Adjusted R-Square | 0.3082 |  | 0.2763 |  |  |  | 0.2749 |  |  |  |
| Sample Size | 263 |  | 292 |  | $565$ |  | . 292 |  | 565 |  |
| Dependent Mean | 1.8281 |  | 1.9307 |  | 2.0270 |  | 1.9307 |  | 2.0270 |  |
| BLACK HCMEN |  |  |  |  |  |  |  |  |  |  |
| variable | 1967 (a) parameter estimate | T-STAT | 1972 (a) Parameter estimate | t-stat | 1983 (a) Parameter ESTIMATE | T-STAT | 1972 (b) Parameter ESTIMATE T-STAT |  | 1983 (b) parameter estimate t-stat |  |
| InTERCEPT | 0.9418 | 1.188 | 0.7013 | 0.582 | 1.7197 | 2.431 | 1.1126 | 0.940 | 1.7759 | 2.528 |
| AGE | -0.0119 | -0.564 | -0.0049 | -0.150 | -0.0331 | -1.782 | -0.0052 | -0.164 | -0.0289 | -1.554 |
| EXPERIENCE | 0.0025 | 0.533 | -0.0033 | -0.451 | 0.0352 | 5.844 | -0.0092 | -1.236 | 0.0267 | 3.561 |
| SCHOOL IHG | 0.0979 | 8.837 | 0.1215 | 7.623 | 0.0819 | 7.377 | 0.1062 | 6.394 | 0.0757 | 6.592 |
| SOUTH | -0.3772 | -5.419 | -0.2608 | -2.687 | -0.1935 | -3.247 | -0.3241 | -3.319 | -0.2132 | -3.551 |
| SMSA | 0.1623 | 2.145 | 0.1817 | 1.894 | 0.1168 | 1.768 | 0.1337 | 1.403 | 0.1349 | 2.038 |
| NO CHILIDREN | 0.1959 | 1.759 | -0.1062 | -0.707 | -0.0635 | -0.736 | -0.1439 | -0.979 | -0.0777 | -0.904 |
| ONE CHILD | -0.0675 | -0.571 | 0.0822 | 0.613 | 0.0239 | 0.318 | 0.1144 | 0.872 | 0.0191 | 0.257 |
| THO CHILDREN | 0.0903 | 1.069 | -0.0490 | -0.418 | -0.1761 | -2.689 | -0.0279 | -0.243 | -0.1578 | -2.402 |
| years out in past five |  |  |  |  |  |  | -0.0887 | -2.582 | -0.0589 | -1.914 |
| Adjusted R-square | 0.5527 |  | 0.4849 |  | 0.4228 |  | 0.5115 |  | 0.4314 |  |
| Sample Size |  |  | 113 |  | 185 |  | 113 |  | 185 |  |
| Dependent Hean | 1.4559 |  | 1.7498 |  | 1.8415 |  | 1.7498 |  | 1.8415 |  |

Table 8-5
Determinants of Log Wages, Ages 40-44

| HON BLACK HOMEN |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| variable | 1967 (a) Parameter ESTiMATE | Y-Stat | 1972 (a) parameter Estimate | T-STAT | 1977 (a) parameter ESTIMATE | f-Stat | 1972 (b) parameter estimate | i-stat | 1977 (b) parameter ESTIMATE | T-Stat |
| [ NTERCEPT | -0.0517 | -0.876 | 2.1714 | 3.647 | 2.0327 | 2.765 | 2.3477 | 3.991 | 1.9586 | 2.728 |
| AGE | 0.0307 | 2.285 | -0.0305 | -2.198 | -0.0355 | -2.031 | -0.0304 | -2.222 | -0.0287 | -1.669 |
| EXPERIEHCE | 0.0119 | 3.900 | 0.0189 | 5.436 | 0.0228 | 5.108 | 0.0133 | 3.497 | 0.0144 | 2.895 |
| SCHOOLING | 0.0668 | 8.727 | 0.0610 | 7.376 | 0.0805 | 7.615 | 0.0567 | 6.869 | 0.0747 | 7.137 |
| SOUTH | -0.0814 | -1.838 | -0.0708 | -1.654 | -0.0911 | -1.778 | -0.0751 | -1.782 | -0.0743 | -1.479 |
| SHSA | 0.1147 | 2.577 | 0.1149 | 2.644 | 0.1125 | 2.156 | 0.1128 | 2.707 | 0.1498 | 2.877 |
| ho CHILDREN | 0.1279 | 2.116 | 0.0017 | 0.024 | 0.0049 | 0.066 | 0.0292 | 0.429 | 0.0482 | 0.655 |
| ONE CHILD | -0.0035 | -0.058 | 0.1098 | 1.365 | 0.0857 | 0.953 | 0.1192 | 1.504 | 0.0891 | 1.015 |
| THO CHILDREM | 0.1240 | 2.552 | 0.0279 | 0.696 | -0.0080 | -0.140 | 0.0253 | 0.567 | 0.0020 | 0.035 |
| years out in past five |  |  |  |  |  |  | -0.0547 | -3.393 | -0.0708 | -3.470 |
| Adjusted R-Square | 0.3231 |  | 0.2575 |  | 0.3134 |  | 0.2807 |  | 0.3453 |  |
| Sample size | 329 |  | 336 |  | 235 |  | 336 |  | 235 |  |
| Dependent Mean | 1.8152 |  | 1.9589 |  | 1.9780 |  | 1.9589 |  | 1.9780 |  |
| BLACK HOMEN |  |  |  |  |  |  |  |  |  |  |
| Variable | 1967 (a) parameter estimate | T-STAT | 1972 (a) PARAMETER EStimate | T-STAT | 1977 (a) PARAMETER ESTIMATE | T-STAT | 1972 (b) Parameter estimate | T-STAT | 1977 (b) parameter ESTIMATE | T-STAT |
| Intercept | 0.7106 | 0.700 | 0.2012 | 0.205 | -1.8349 | -1.451 | 0.2041 | 0.206 | -1.9695 | -1.538 |
| AGE | -0.0116 | -0.490 | 0.0140 | 0.614 | 0.0704 | 2.338 | 0.0138 | 0.597 | 0.0722 | 2.382 |
| EXPERIENCE | 0.0151 | 3.188 | 0.0007 | 0.151 | -0.0142 | -2.159 | 0.0008 | 0.157 | -0.0136 | -2.081 |
| SCHOOLING | 0.0970 | 9.508 | 0.0874 | 7.985 | 0.0702 | 4.644 | 0.0875 | 7.730 | 0.0728 | 4.679 |
| SOUTH | -0.2783 | -3.629 | -0.3346 | -4.547 | -0.0599 | -0.624 | -0.3350 | -4.516 | -0.0479 | -0.491 |
| SHSA | 0.1305 | 1.653 | 0.2537 | 3.439 | 0.3330 | 3.677 | 0.2539 | 3.425 | 0.3427 | 3.735 |
| ho children | 0.0678 | 0.731 | 0.2011 | 1.752 | -0.1349 | -0.912 | 0.2014 | 1.746 | -0.1347 | -0.908 |
| ONE CHILD | -0.0250 | -0.298 | -0.1811 | -1.378 | 0.0711 | 0.518 | -0.1810 | -1.372 | 0.0646 | 0.468 |
| THO CHILDREN | 0.2883 | 2.963 | 0.0825 | 0.970 | 0.0611 | 0.593 | 0.0826 | 0.966 | 0.0511 | 0.491 |
| years out in past five |  |  |  |  |  |  | 0.0015 | 0.053 | 0.0284 | $0.738^{\prime}$ |
| Adjusted R-square | 0.5141 |  | 0.5104 |  | 0.4484 |  | 0.5065 |  | 0.4453 . |  |
| Sample size | 152 |  | 136 |  | 91 |  | 136 |  | 91 |  |
| Dependent Mean | 1.4533 |  | 1.7127 |  | 1.8804 |  | 1.7127 |  | 1.8804 |  |


| NON BLACK WONEH |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Variable | 1972 <br> Parameter estimate | (a) T-STAT | PARAMETER ESTIMATE | $7 \text { (a) }$ T-STAT | 1982 (a) <br> parameter <br> estimate | T-STAT | 1972 (b) <br> parameter <br> estimate | T-STAT | 197 <br> PARAMETER ESTIMATE | (b) T-STAT | 1982 (b) <br> PARANETER <br> ESTIMATE | T-StAT |
| IHTERCEPT | 0.9644 | 1.534 | 1.7803 | 2.460 | 2.8926 | 3.239 | 1.2359 | 1.892 | 1.7942 | 2.482 | 3.0 .3 | 3.485 |
| AGE | -0.0041 | -0.308 | -0.0217 | -1.426 | -0.0549 | -2.886 | -0.0061 | -0.470 | -0.0208 | -1.368 | -0.0540 | 2.890 |
| EXPERIENCE | 0.0151 | 5.477 | 0.0200 | 5.466 | 0.0244 | 5.530 | 0.0109 | 3.695 | 0.0183 | 4.708 | 0.0177 | 3.686 |
| SCHOOLING | 0.0751 | 9.533 | 0.0634 | 6.968 | 0.0876 | 7.521 | 0.0702 | 8.897 | 0.0618 | 6.743 | 0.0830 | 7.206 |
| SOUTH | -0.0814 | -1.918 | -0.0705 | -1.504 | -0.0331 | -0.591 | -0.0884 | -2.114 | -0.0701 | -1.497 | -0.0314 | -0.572 |
| SHSA | 0.0596 | 1.431 | 0.1338 | 2.856 | 0.1830 | 3.175 | 0.0631 | 1.538 | 0.1364 | 2.911 | 0.1974 | 3.479 |
| HO CHILDREN | 0.0702 | 1.095 | 0.1245 | 1.671 | -0.1614 | -1.754 | 0.0883 | 1.395 | 0.1318 | $1.76{ }^{\prime}$ | -0.1226 | -1.346 |
| ONE CHILD | 0.0212 | 0.360 | 0.0485 | 0.614 | -0.0909 | -0.929 | 0.0125 | 0.216 | 0.0532 | 0.673 | -0.0596 | -0.618 |
| THO CHILDREN | 0.0325 | 0.694 | 0.0425 | 0.831 | 0.0446 | 0.672 | 0.0407 | 0.881 | 0.0426 | 0.833 | 0.0551 | 0.844 |
| YEARS OUT IL PASt five |  |  |  |  |  |  | -0.0538 | -3.523 | -0.0362 | -1.247 | -0.0855 | -3.279 |
| Adjusted R-Square | 0.3264 |  | 0.3152 |  | 0.2730 |  | 0.3473 |  | 0.3168 |  | 0.3001 |  |
| Sample Size | 367 |  | 250 |  | 261 |  | 367 |  | 250 |  | 261 |  |
| Dependent Mean | 1.9365 |  | 2.0259 |  | 1.9938 |  | 1.9365 |  | 2.0259 | ! | 1.9938 |  |
| BLACK HOWEN |  |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  | $1977 \text { (a) }$ <br> PARAMETER |  | $1982 \text { (a) }$ <br> PARALETER |  | 1972 (b) parañeter |  | $1977 \text { (b) }$ |  | $1982 \text { (b) }$ |  |
| VARIABLE | ESTIMATE | T-STAT | ESTIMATE T-STAT |  | ESTIMATE | T-STAT | ESTIMATE T-STAT |  | ESTIMATE | T-STAT | ESTIMATE | T-STAT |
| INTERCEPT | 0.7685 | ค 584 | 0.0083 | 0.008 | -0.5373 | -0.339 | 0.4251 | 0.321 | 0.0784 | 0.071 | -0.8619 | -0.571 |
| AGE | -0.0021 | 0.075 | J. 3167 | 0.744 | 0.0344 | 1.016 | 0.0007 | 0.024 | 0.0163 | 0.722 | 0.0470 | 1.444 |
| EXPERIENCE | 0.0048 | 0.975 | 2.0016 | 0.346 | -0.0032 | -0. $5 \cdot \mathrm{P}$ | 0.0088 | 1.604 | 0.0005 | 0.109 | -0.0114 | -1.645 |
| SCHOOL ING | 0.0964 | 8.627 | 1045 | 9.486 | 0.0806 | 4. | 0.1040 | 8.579 | 0.1022 | 8.453 | 0.0772 | 4.498 |
| SOUTH | -0.3126 | -3.530 | - 3829 | -5.002 | -0.2153 | -2.061 | -0.3120 | -3.544 | -0.3850 | -5.002 | -0.2327 | -2.337 |
| SMSA | 0.1652 | 1.769 | 0.1382 | 1.890 | 0.2007 | 1.927 | 0.1716 | 1.846 | 0.1372 | 1.869 | 0.2265 | 2.276 |
| NO CHILDREN | -0.0482 | -0.407 | -0.00 | -0.069 | 0.0652 | 0.413 | -0.0557 | -0.473 | 0.0012 | 0.011 | 0.0340 | 0.226 |
| ONE CHILD | -0.0839 | -0.889 | -0.010s | -0.082 | -0.0597 | -0.*) | -0.0823 | -0.877 | 0.0094 | 0.069 | -0.0587 | -0.361 |
| THO CHILOREN | 0.0719 | 0.648 | 0.0096 | 0.115 | -0.1622 | -1.35 | 0.0758 | 0.686 | 0.0217 | 0.248 | -0.1938 | -1.698 |
| Years out in past five |  |  |  |  |  |  | 0.0624 | 1.566 | -0.0170 | -0.470 | -0.3006 | -2.988 |
| Adjusted R-Square | 0.4538 |  | 0.6001 |  | 0.3822 |  | 0.4599 |  | 0.5970 |  |  |  |
| Sample size | 137 |  | 109 |  | $85$ |  | 137 |  | 109 |  | 85 |  |
| Deperdent Mean | 1.6938 |  | 1.8274 |  | 1.9101 |  | 1.6938 |  | 1.8274 |  | 1.9101 |  |

Table B-7
Determinants of Log Hages, Ages 50-54



[^0]:    ${ }^{1}$ Female-male hourly earnings ratios among white full-time workers have changes as follows, by age, from 1977 to 1988: ages 25-34, from 69.4 to 80.3 percent; ages $35-44$, from 56.2 to 66.9 percent; ages $45-54$, from 54 to 62.8 percent; ages $55-64,57.8$ to 60.8 percent. (U.S. Bureau of the Census, Current Population Survey, March Supplements Public Use Tapes).

[^1]:    ${ }^{2}$ The pioneering work in this area is that of Mincer (1962).
    ${ }^{3}$ See also Kiliingsworth (1983) for a comprehensive survey of empirical results for female labor supply.

[^2]:    ${ }^{4}$ By reading down the diagonal one can examine the change in labor force participation rates for a cohort as it ages. For example, the cohort ages 3034 in 1967 was $35-39$ in 1972, 40-44 in 1977, and so on.

[^3]:    ${ }^{6}$ See also work on dynamic labor supply, e.g., Heckman and MaCurdy (1980) and Flinn and Heckman (1982, 1983).

[^4]:    ${ }^{7}$ See Killingsworth (1983), Killingsworth and Heckman (1986), and MaCurdy (1985) for reviews of dynamic models and empirical results and see Heckman and Singer (1986) for a brief outline of a similar dynamic labor supply model.
    ${ }^{8}$ This index is analogous to the one-period decision rule common to the labor force participation models of Heckman (1974, 1977), and Cogan (1977), among others.

[^5]:    ${ }^{9}$ See Yi, et al. (1987), Flinn and Heckman (1982, 1983), and Heckman and singer (1986) for degcriptiong of the multiple-spell hazard rate model and estimation.
    ${ }^{10}$ The NLS Young Women's sample began skipping years in the late 1970 s so a full sequence of annual observations is not available. However, through the use of retrospective questions, it is possible to fill in important information for the misaing years. Given the nature of the retrospective questions, we define a work year as one in which the respondent worked at least six months to link the retrospective and prospective data. Unfortunately this definition combines together the states "unemployment" and "out of the labor force" as "non-work."

[^6]:    ${ }^{11}$ Experience is measured as years since schooling was completed for the 20 to 24 year olds and as years since age 18 for the 15 to 19 year olds.
    ${ }^{12}$ For the 15 to 19 year olds, there was so little variation in the number of children older than 6 that the multiple spell models were not estimable when they included that variable.
    ${ }^{13}$ Missing survey years ruled out including the unemployment rate in the local labor market as an additional explanatory variable.

[^7]:    ${ }^{14}$ To adjust for the oversampling of Hispanics in the NLS-Y, we drew a random sample of all Hispanic women such that our sample proportion corresponds with the age and sex specific population proportion as reported in the 1980 Census.

[^8]:    ${ }^{15}$ We use a quadratic duration specification. Likelihood ratio tests indicated that we could reject both Weibull and Gompertz specifications in favor of the quadratic model.
    ${ }^{16}$ Early results baged on a non-parametric mixture distribution with two support points indicated little change in any of the estimated parameters.

[^9]:    ${ }^{17}$ One topic that remaing for the future is the potential endogeneity of some explanatory variables (measures of fertility in particular). A recent paper by Olsen and Farkag (1989) outlines one methodological approach to treating endogenous variables within a hazard rate model.

